SUPPLEMENT TO "UNWILLING TO TRAIN?—FIRM RESPONSES TO THE COLOMBIAN APPRENTICESHIP REGULATION" (Econometrica, Vol. 90, No. 2, March 2022, 507-550)<br>SANTiago Caicedo<br>John E. Walker Department of Economics, Clemson University<br>Miguel Espinosa<br>Department of Economics, Universitat Pompeu Fabra and Management and Technology Department, Bocconi University<br>Arthur Seibold<br>Department of Economics, University of Mannheim

This supplement contains the Online Appendices of "Unwilling to Train?-Firm Responses to the Colombian Apprenticeship Regulation" by Caicedo, Espinosa, and Seibold (2022). Appendix A shows additional tables and figures. Appendix B provides additional information on training courses and apprentices. Appendix C presents reduced-form results using an alternative sector classification. Appendix D contains model proofs and extensions. Finally, Appendix E provides details of the quantitative exercises.

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## APPENDIX A: AdDitional Figures and Tables



Figure A1.-Firm Size Distribution Pre and Post-Reform (Selected Years). Notes: The figure shows the distribution of full-time workers in 1995 (first available year), 2002 (last year before the reform), 2003 (first year after the reform), and 2009 (last available year), using a bin size of one. Vertical dashed lines denote the regulation thresholds. Graphs for all other years are similar and available upon request.


Figure A2.-Apprentices Prereform, Fees, and Fines. Notes: Panel (a) of the figure shows the average number of apprentices by firm size bin in high-skill and low-skill sector firms, prereform (1995-2002). The black dashed lines show the minimum and maximum apprentice quotas. Panels (b) and (c) show the fraction of firms paying fees and fines by sector in the post-reform years 2003 to 2009. In all panels, vertical dashed lines denote the regulation thresholds.


Figure A3.-Distribution of Profit Changes. Notes: The figure shows the distribution of percentage changes in firm profits by sector in partial equilibrium (panel a), general equilibrium and the dynamic scenario (both panel b).


Figure A4.-Reduced-Form Effects of the Regulation on Firm Outcomes. Notes: The figure shows yearly difference-in-difference coefficients for the period 1999 to 2006, using 2002 as the base year. The vertical bars denote $95 \%$ confidence intervals based on standard errors clustered at the firm level.


Figure A5.-Reduced-Form Coefficients Versus Model Prediction. Notes: The figure shows a comparison of reduced-form effects versus model-based effects as described in Section 5.5. The hollow diamonds depict the difference-in-difference coefficients from Table VII. The circles show analogously simulated differential effects on firms above vs. below regulation thresholds from the partial-equilibrium model. Vertical bars denote $95 \%$ confidence intervals for both sets of coefficients.
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| TABLE AI Continued. |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | (1) <br> Wood <br> Products | (2) <br> Textile | (3) <br> Food/ Beverage | (4) <br> Mineral NonMetallic | (5) <br> Other Manufacturing | (6) <br> Paper/ Editorial | (7) <br> Metallic Products | (8) <br> Machinery/ Equipment | (9) <br> Chemical Products | Rank Correlation With Baseline Proxy |
| Years of | Value | 7.89 | 8.28 | 8.04 | 7.63 | 8.08 | 10.56 | 8.48 | 9.97 | 10.32 |  |
| Education | Rank | 8 | 5 | 7 | 9 | 6 | 1 | 4 | 3 | 2 | 0.73 |
| (prod./admin) | High/Low | L | (H) | L | L | L | H | H | H | H |  |
| Fraction with | Value | 0.105 | 0.074 | 0.107 | 0.139 | 0.091 | 0.207 | 0.090 | 0.183 | 0.220 |  |
| Tertiary Educ. | Rank | 6 | 9 | 5 | 4 | 7 | 2 | 8 | 3 | 1 | 0.57 |
| (prod./admin) | High/Low | L | L | (H) | H | L | H | L | H | H |  |
| Fraction above | Value | 0.39 | 0.45 | 0.49 | 0.39 | 0.32 | 0.65 | 0.54 | 0.70 | 0.69 |  |
| Median Wage | Rank | 8 | 6 | 5 | 7 | 9 | 3 | 4 | 1 | 2 | 0.75 |
| (prod./admin) | High/Low | L | L | (H) | L | L | H | H | H | H |  |
| Observations (prod./admin) |  | 876 | 16,212 | 11,608 | 5136 | 5103 | 2071 | 3415 | 1842 | 3544 |  |

Note: The table shows different proxies for skill requirements by 2-digit industries. For each characteristic, the first row displays the average value among firms/workers in an industry, the second row shows the implied rank among the nine 2 -digit industries, and the third row indicates whether the industry is classified as "high-skill" or "low-skill" according to this characteristic, where "L" denotes low-skill, "H" denotes high-skill, and "(H)" denotes the marginal high-skill rank 5 out of 9 . Panel A shows our baseline skills proxy, the fraction of professionals, and Panel B shows rankings based on wages observed in the main firm-level data. Panel C displays rankings based on various measures from the Colombian Household Survey (ECH), corresponding to characteristics of workers by their reported industry of employment. All values shown in the table are calculated in the available pre-reform years only, namely 1995 to 2002 in Panels A and B, and 2001 to 2002 in Panel C. The last column shows the Spearman rank correlation of each characteristic with the baseline skills proxy from Panel A.

TABLE AII
SUMMARY STATISTICS PREREFORM.

|  | (1) | (2) | (3) | (4) |
| :---: | :---: | :---: | :---: | :---: |
|  | All Sectors | Low-Skill | High-Skill | t-test High vs. Low-Skill |
| Apprentices | $\begin{gathered} 0.18 \\ (0.60) \end{gathered}$ | $\begin{gathered} 0.17 \\ (0.59) \end{gathered}$ | $\begin{gathered} 0.19 \\ (0.60) \end{gathered}$ | 0.014 |
| Workers | $\begin{gathered} 55.14 \\ (105.26) \end{gathered}$ | $\begin{gathered} 55.50 \\ (112.80) \end{gathered}$ | $\begin{gathered} 54.79 \\ (97.21) \end{gathered}$ | 0.419 |
| Workers (Survey) | $\begin{gathered} 61.03 \\ (142.59) \end{gathered}$ | $\begin{gathered} 62.33 \\ (153.32) \end{gathered}$ | $\begin{gathered} 59.74 \\ (131.07) \end{gathered}$ | 0.029 |
| Fraction professionals | $\begin{gathered} 0.07 \\ (0.12) \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.11) \end{gathered}$ | $\begin{gathered} 0.08 \\ (0.13) \end{gathered}$ | 0.000 |
| Fraction admin workers | $\begin{gathered} 0.36 \\ (0.25) \end{gathered}$ | $\begin{gathered} 0.37 \\ (0.27) \end{gathered}$ | $\begin{gathered} 0.35 \\ (0.23) \end{gathered}$ | 0.000 |
| Fraction production workers | $\begin{gathered} 0.57 \\ (0.27) \end{gathered}$ | $\begin{gathered} 0.58 \\ (0.28) \end{gathered}$ | $\begin{gathered} 0.57 \\ (0.25) \end{gathered}$ | 0.521 |
| Output | $\begin{gathered} 9,850,148 \\ (26,403,640) \end{gathered}$ | $\begin{gathered} 9,325,289 \\ (24,970,662) \end{gathered}$ | $\begin{gathered} 10,371,309 \\ (27,744,018) \end{gathered}$ | 0.000 |
| Value added | $\begin{gathered} 4,383,855 \\ (12,330,128) \end{gathered}$ | $\begin{gathered} 3,910,267 \\ (11,378,001) \end{gathered}$ | $\begin{gathered} 4,850,617 \\ (13,185,101) \end{gathered}$ | 0.000 |
| Profits | $\begin{gathered} 2,809,289 \\ (8,899,159) \end{gathered}$ | $\begin{gathered} 2,505,919 \\ (8,176,656) \end{gathered}$ | $\begin{gathered} 3,108,285 \\ (9,548,693) \end{gathered}$ | 0.000 |
| Wage bill (permanent workers) | $\begin{gathered} 1,283,054 \\ (3,147,877) \end{gathered}$ | $\begin{gathered} 1,134,166 \\ (2,897,829) \end{gathered}$ | $\begin{gathered} 1,430,893 \\ (3,371,448) \end{gathered}$ | 0.000 |
| Total wage bill | $\begin{gathered} 1,537,701 \\ (3,684,091) \end{gathered}$ | $\begin{gathered} 1,352,669 \\ (3,315,540) \end{gathered}$ | $\begin{gathered} 1,721,430 \\ (4,008,340) \end{gathered}$ | 0.000 |
| Wage per worker (permanent workers) | $\begin{gathered} 17,204 \\ (15,239) \end{gathered}$ | $\begin{gathered} 15,831 \\ (14,841) \end{gathered}$ | $\begin{gathered} 18,524 \\ (15,497) \end{gathered}$ | 0.000 |
| Capital/Output | $\begin{gathered} 0.69 \\ (0.91) \end{gathered}$ | $\begin{gathered} 0.66 \\ (0.94) \end{gathered}$ | $\begin{gathered} 0.72 \\ (0.88) \end{gathered}$ | 0.000 |
| Intermediates/Output | $\begin{gathered} 0.54 \\ (0.18) \end{gathered}$ | $\begin{gathered} 0.55 \\ (0.19) \end{gathered}$ | $\begin{gathered} 0.52 \\ (0.17) \end{gathered}$ | 0.000 |
| Observations | 57,694 | 28,745 | 28,949 |  |
| Firms | 10,244 | 5465 | 5388 |  |

Note: The table shows summary statistics for the full sample in the pre-reform years 1995 to 2002. All monetary variables are in 2009 thousands of pesos. In columns (1) to (3), standard deviations are in parentheses. Column (4) shows the p-value of a $t$-test of equality of means in high- versus low-skill sectors.

TABLE AIII
CORRELATES OF RESPONSES TO REGULATION.

|  | $(1)$ <br> Choose Maximum <br> Quota | $(2)$ <br> Choose Mininimum <br> Quota | $(3)$ <br> Between Minimum <br> and Maximum | $(4)$ <br> Pay Fee to <br> Avoid Apprentices |
| :--- | :---: | :---: | :---: | :---: |
|  | -0.63 | 0.047 | -0.044 | 0.57 |
| High-skill Sector | $(0.0099)$ | $(0.011)$ | $(0.0034)$ | $(0.0049)$ |
| Number of workers | 0.00021 | -0.00017 | -0.000019 | -0.0000072 |
|  | $(0.000060)$ | $(0.000056)$ | $(0.000030)$ | $(0.000025)$ |
| Wage per worker | 0.0016 | -0.0015 | -0.000055 | 0.00012 |
|  | $(0.00045)$ | $(0.00050)$ | $(0.00014)$ | $(0.00021)$ |
| Log output | -0.016 | -0.00056 | 0.016 | -0.0011 |
|  | $(0.0063)$ | $(0.0065)$ | $(0.0026)$ | $(0.0034)$ |
| Output per worker | -0.00012 | 0.00014 | -0.000028 | -0.000020 |
|  | $(0.000030)$ | $(0.000032)$ | $(0.0000087)$ | $(0.000015)$ |
| Profit rate | -0.17 | 0.24 | -0.043 | -0.037 |
|  | $(0.047)$ | $(0.052)$ | $(0.016)$ | $(0.027)$ |
| Capital/Output | -0.056 | 0.051 | 0.0037 | -0.0062 |
|  | $(0.0079)$ | $(0.0083)$ | $(0.0028)$ | $(0.0038)$ |
| Intermediates/Output | -0.21 | 0.26 | -0.055 | 0.0055 |
|  | $(0.044)$ | $(0.048)$ | $(0.017)$ | $(0.023)$ |
| Mean dep. var. | 0.29 | 0.30 | 0.053 | 0.32 |
| Observations | 21265 | 21265 | 21265 | 21265 |
| R-squared | 0.48 | 0.018 | 0.016 | 0.38 |

Note: The table shows regressions of indicators for different types of responses to the apprenticeship regulation on pre-reform firm characteristics, where the characteristics are averaged across prereform years by firm. All monetary variables in 2009 pesos, in units of thousands. All regressions include year fixed effects. Standard errors clustered by firm in parentheses.

## TABLE AIV

CORRELATES OF BUNCHING BEHAVIOR.

|  | Bunchers Above | Bunchers Below | All Firms |
| :--- | :---: | :---: | ---: |
| Fraction of all firms | 0.04 | 0.08 | 1.00 |
| Share in high-skill sector | 0.07 | 0.88 | 0.56 |
| Mean number of workers | 61.04 | 51.14 | 42.19 |
| Choose maximum quota | 0.72 | 0.18 | 0.21 |
| Choose minimum quota | 0.17 | 0.49 | 0.56 |
| Between min./max. | 0.06 | 0.02 | 0.03 |
| Pay fee to avoid apprentices | 0.05 | 0.27 | 0.17 |
| Observations | 2167 | 4154 | 50,691 |
| Firms | 1468 | 2624 | 10,740 |

[^1]TABLE AV
REDUCED-FORM EFFECTS OF THE POLICY ON ADDITIONAL FIRM OUTCOMES.

|  | (1) | (2) | (3) | (4) <br> (5) <br> Panel A: Workers by Type |  |  | (7) | (8) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Workers |  | Professionals |  | Admin Workers |  | Production Workers |  |
|  | High-Skill | Low-Skill | High-Skill | Low-Skill | High-Skill | Low-Skill | High-Skill | Low-Skill |
| Above*Post | $\begin{aligned} & -0.855 \\ & (0.863) \end{aligned}$ | $\begin{aligned} & -1.781 \\ & (1.036) \end{aligned}$ | $\begin{gathered} -0.0556 \\ (0.224) \end{gathered}$ | $\begin{gathered} -0.0649 \\ (0.185) \end{gathered}$ | $\begin{aligned} & -0.199 \\ & (0.359) \end{aligned}$ | $\begin{aligned} & -0.956 \\ & (0.445) \end{aligned}$ | $\begin{aligned} & -0.813 \\ & (0.652) \end{aligned}$ | $\begin{aligned} & -0.898 \\ & (0.781) \end{aligned}$ |
| Mean (Prereform) | 30.46 | 30.30 | 2.581 | 1.500 | 9.393 | 10.08 | 19.10 | 19.08 |
| Observations | 8491 | 6357 | 7471 | 5587 | 8491 | 6357 | 8491 | 6357 |
| R-squared | 0.904 | 0.894 | 0.767 | 0.867 | 0.887 | 0.859 | 0.874 | 0.883 |

Panel B: Other Measures of Labor Input
Workers Workers (Survey) Temp. Workers Outsourced Workers High-Skill Low-Skill High-Skill Low-Skill High-Skill Low-Skill High-Skill Low-Skill

| Above*Post | -0.855 | -1.781 | -0.757 | -2.372 | 0.604 | 0.392 | -0.131 | 2.134 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $(0.863)$ | $(1.036)$ | $(0.891)$ | $(1.304)$ | $(0.643)$ | $(1.805)$ | $(1.137)$ | $(2.247)$ |
| Mean (Prereform) | 30.46 | 30.30 | 31.46 | 32.33 | 3.329 | 4.831 | 4.719 | 7.299 |
| Observations | 8491 | 6357 | 8491 | 6357 | 8491 | 6357 | 8491 | 6357 |
| R-squared | 0.904 | 0.894 | 0.907 | 0.899 | 0.687 | 0.597 | 0.752 | 0.880 |

## Panel C: Other Inputs

| Log Capital |  | Log Intermediates |  | Log Energy |  | Log Water |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| High-Skill | Low-Skill | High-Skill | Low-Skill | High-Skill | Low-Skill | High-Skill | Low-Skill |
| $\begin{aligned} & 0.00349 \\ & (0.0483) \end{aligned}$ | $\begin{gathered} 0.0272 \\ (0.0689) \end{gathered}$ | $\begin{gathered} -0.0268 \\ (0.0382) \end{gathered}$ | $\begin{gathered} -0.00554 \\ (0.0502) \end{gathered}$ | $\begin{gathered} -0.0228 \\ (0.0433) \end{gathered}$ | $\begin{gathered} -0.0112 \\ (0.0497) \end{gathered}$ | $\begin{aligned} & -0.0304 \\ & (0.0562) \end{aligned}$ | $\begin{gathered} 0.0298 \\ (0.0760) \end{gathered}$ |
| 13.44 | 13.34 | 13.58 | 13.82 | 11.67 | 11.98 | 9.978 | 9.837 |
| 8491 | 6357 | 8491 | 6357 | 8491 | 6357 | 8407 | 6288 |
| 0.834 | 0.850 | 0.861 | 0.871 | 0.876 | 0.882 | 0.787 | 0.753 |

Panel D: Output, Value Added, Productivity
Log Output Log Value Added Log Output per Worker TFP
High-Skill Low-Skill High-Skill Low-Skill High-Skill Low-Skill High-Skill Low-Skill

| Above*Post | -0.0302 | -0.00697 | -0.0429 | 0.0190 | 0.0139 | 0.0122 | -0.0116 | -0.0128 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $(0.0355)$ | $(0.0472)$ | $(0.0423)$ | $(0.0553)$ | $(0.0331)$ | $(0.0465)$ | $(0.0284)$ | $(0.0361)$ |
| Mean (Prereform) | 14.31 | 14.46 | 13.52 | 13.49 | 11.21 | 11.35 | 5.509 | 5.581 |
| Observations | 8491 | 6357 | 8491 | 6357 | 8376 | 6211 | 8058 | 5909 |
| R-squared | 0.866 | 0.862 | 0.814 | 0.786 | 0.772 | 0.813 | 0.794 | 0.822 |

Note: The table shows results from difference-in-difference regressions as described by equation (4). Regressions are run on the threshold sample, using years 1999 to 2006, and include year and firm fixed effects. Standard errors clustered at the firm level in parentheses.

## APPENDIX B: Additional Information on Training Courses and Apprentices

The results in this Appendix are based on several additional data sources. Two data sets allow us to gather additional information on apprentices' characteristics. The first data set is a survey of school-to-work transitions (ETET) focused on young individuals aged between 14 and 29 years. The survey was conducted by the Colombian government in 2013 and 2015. The second data set are monthly administrative records from the social security system (PILA) gathered by the Colombian Ministry of Health and Social Protection and processed by the Colombian National Statistical Agency (DANE). This data is available for 2015 and 2016. In both data sets, we pool the available years. Moreover, we use the Colombian household survey (ECH), which is available for the years 2001 to 2006. This data set is administered by DANE and constitutes the first statistical tool used by the Colombian government to study labor markets. Unfortunately, apprentices cannot be identified in the ECH data, but we can observe other workers' characteristics.

This Appendix is organized as follows. Tables BI and BII show information on specific apprenticeship training courses. Table BIII shows characteristics of apprentices and other workers based on various data sources. Figures B1 and B2 as well as Table BIV focus on apprentices in the PILA data who complete their apprenticeship during the sample period. Specifically, Figures B1 and B2 show wages of this group around the month of graduation, compared to a control group of apprentices that do not graduate throughout this period. Table BIV shows transition probabilities between firms and sectors and wages of apprentices after they complete their apprenticeship. Finally, Table BV shows the occupational distribution by sector as observed in the ECH data.

TABLE BI
LARGEST TRAINING COURSES BY ENROLLMENT (ALL SECTORS, 2015).

| Sector | Training Course | Trainees Enrolled |
| :--- | :--- | :---: |
| General | Systems | 2870 |
| Services | Administrative Assistance | 2433 |
| Services | Accounting of Commercial and Financial Operations | 2121 |
| Health | Environment Management | 1187 |
| Agricultural | Agricultural Production | 1057 |
| Health | Occupational Safety | 988 |
| Services | Accounting and Finances | 902 |
| General | Analysis and Development of Information Systems | 854 |
| Commercial | Sales of Products and Services | 796 |
| Services | Administrative Management | 746 |
| Commercial | Business Management | 737 |
| Services | Kitchen | 643 |
| Electricity | Electrical Residential Installations | 597 |
| General | Software Programming | 555 |
| Electricity | Maintenance of Computer Equipment | 514 |
| Agricultural | Ecological Agricultural Systems | 512 |
| General | Human Resources | 508 |
| Construction | Building Construction | 487 |
| Services | Assistance for Organization of Archives | 448 |
| Commercial | Market Management | 445 |
| Services | Food Agroindustry | 441 |

Note: The table shows the 20 largest training courses by enrollment in SENA in 2015. "General" refers to the training course that span several sectors (denoted as "transversal" in original data). Apprentices cannot be enrolled in more than one training course simultaneously. Source: SENA (2016).
TABLE BII
Examples of specific training courses (manufacturing, 2020).

| Sector | Training Course | Required Education (Years) | Length of Training (Months) |
| :---: | :---: | :---: | :---: |
| Manufacturing | Setting up Analysis and Testing Laboratories for Industry | 9 | 15 |
| Manufacturing | Inspection and Testing with Nondestructive Processes | 9 | 15 |
| Manufacturing | Industrial Manufacture of Outerwear | 5 | 6 |
| Manufacturing | Maintenance of Industrial Automation | 9 | 15 |
| Manufacturing | Paint Coating for Transportation Equipment | 9 | 15 |
| Manufacturing | Chemical Industry Processes | 11 | 27 |
| Manufacturing | Metal Joinery | 9 | 15 |
| Manufacturing | Light Vehicle Maintenance | 9 | 15 |
| Manufacturing | Maintenance of Diesel Engines | 9 | 15 |
| Manufacturing | Industrial Production Management | 11 | 24 |
| Manufacturing | Mechatronic Automotive Maintenance | 11 | 24 |
| Manufacturing | Cabinet Making | 5 | 15 |
| Manufacturing | Manufacture of Products with Fiberglass and Composite Materials | 9 | 15 |
| Manufacturing | Cutting of Footwear and Leather Goods | 5 | 9 |
| Manufacturing | Polymer Transformation by Injection | 9 | 15 |
| Manufacturing | Maintenance of Motorcycles and Motorcars | 9 | 15 |
| Manufacturing | Carpentry | 5 | 15 |
| Manufacturing | Food Processing | 11 | 24 |
| Manufacturing | Aluminum Joinery | 9 | 12 |
| Manufacturing | Machining on a Conventional Lathe and Milling Machine | 9 | 15 |
| Manufacturing | Industrial Machinery Mechanics | 9 | 15 |
| Manufacturing | Armed Jewelry | 9 | 15 |
| Manufacturing | Electrical Maintenance and Electronic Control of Automobiles | 9 | 15 |



TABLE BIII
Characteristics of apprentices.

|  | (1) <br> Apprentices <br> (ETET) | (2) Other Young Workers (ETET) | (3) <br> Apprentices (PILA) | (4) Other Workers (PILA) | (5) <br> Apprentices <br> (EAM) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Fraction Female | 68.4\% | 44.7\% | 57.9\% | 39.2\% | 54.7\% |
| Age | 20.94 | 23.35 | 23.51 | 35.18 |  |
| Fraction 14 to 19 | 35.7\% | 18.4\% | 19.5\% | 2.5\% |  |
| Fraction 20 to 24 | 51.0\% | 39.1\% | 50.8\% | 17.3\% |  |
| Fraction 25 to 29 | 13.3\% | 42.5\% | 15.3\% | 20.5\% |  |
| Education |  |  |  |  |  |
| Primary or less | 0.0\% | 5.4\% |  |  |  |
| Some Secondary | 2.0\% | 14.0\% |  |  |  |
| High School | 18.4\% | 32.7\% |  |  |  |
| Technical/vocational | 49.0\% | 24.6\% |  |  |  |
| College | 30.6\% | 23.3\% |  |  |  |
| Last monthly wage | 513,367 | 808,493 |  |  |  |
| Wage relative to minimum wage |  |  | 1.04 | 2.43 |  |
| Socioeconomic status |  |  |  |  |  |
| Low | 44.8\% | 66.2\% |  |  |  |
| Medium | 51.0\% | 31.2\% |  |  |  |
| High | 4.2\% | 2.6\% |  |  |  |
| Occupational group |  |  |  |  |  |
| Professional |  |  |  |  | 22.2\% |
| Production Worker |  |  |  |  | 39.5\% |
| Administrative worker |  |  |  |  | 38.2\% |
| Field of training/occupation |  |  |  |  |  |
| Humanities and arts | 4.1\% | 5.8\% |  |  |  |
| Social sciences and business | 35.7\% | 24.3\% |  |  |  |
| Sciences | 3.1\% | 3.6\% |  |  |  |
| Engineering, industry and construction | 23.5\% | 27.5\% |  |  |  |
| Agriculture | 0.0\% | 1.3\% |  |  |  |
| Health and social services | 13.3\% | 17.0\% |  |  |  |
| Education | 1.0\% | 6.6\% |  |  |  |
| General services | 9.2\% | 5.3\% |  |  |  |
| Other or not specified | 10.2\% | 8.6\% |  |  |  |
| Sector of training/employment |  |  |  |  |  |
| Agriculture | 0.0\% | 0.7\% |  |  |  |
| Mining | 0.0\% | 0.3\% |  |  |  |
| Manufacturing | 24.5\% | 12.9\% |  |  |  |
| Electricity | 3.1\% | 0.5\% |  |  |  |
| Construction | 2.0\% | 6.3\% |  |  |  |
| Trade and consumer services | 22.4\% | 35.5\% |  |  |  |
| Transport, storage, and communication | 2.0\% | 9.7\% |  |  |  |
| Finance | 6.1\% | 2.5\% |  |  |  |
| Real State | 9.2\% | 9.5\% |  |  |  |
| Public administration | 28.6\% | 13.3\% |  |  |  |
| Other services | 2.0\% | 8.8\% |  |  |  |

TABLE BIII
Continued.
$\left.\begin{array}{lccccc}\hline & \text { (1) } & \begin{array}{c}\text { (2) } \\ \text { Other } \\ \text { Apprentices } \\ \text { (ETET) }\end{array} & \begin{array}{c}\text { (3) } \\ \text { Young Workers } \\ \text { (ETET) }\end{array} & \begin{array}{c}\text { (4) } \\ \text { Apprentices } \\ \text { (PILA) }\end{array} & \begin{array}{c}\text { Other } \\ \text { Workers } \\ \text { (PILA) }\end{array}\end{array} \begin{array}{c}\text { (5) } \\ \text { Apprentices } \\ \text { (EAM) }\end{array}\right]$

Note: The table summarizes characteristics of apprentices and workers based on various data sources. Columns (1) and (2) shows individual characteristics from the ETET data. "Other young workers" are 14 to 29 years old. Last monthly wage in Colombian Pesos. Columns (3) and (4) show individual characteristics from the PILA data. Column (5) shows additional information from the firm-level EAM data, where the shares of apprentices (including interns) by gender and by broad occupational group are observed for a subset of firms and years. Number of observations corresponds to number of individuals in Columns (1) to (4) and to number of firms in Column (5).

TABLE BIV
FIRM AND SECTOR TRANSITIONS AFTER COMPLETING THE APPRENTICESHIP.

## (1)

(2)
(3)

Trained in
All High-Skill Sector Low-Skill Sector
Panel A: Probability to Remain in PILA Data after Training
Remain in Data

| Probability | $23.94 \%$ | $23.23 \%$ | $24.48 \%$ |
| ---: | ---: | ---: | ---: |
| Observations | 89,302 | 38,814 | 50,488 |

Panel B: Firm and Sector Transitions
Stay in Same Firm

| Probability conditional on remaining in data | $74.80 \%$ | $74.39 \%$ | $75.10 \%$ |
| :--- | :---: | :---: | :---: |
| Unconditional Probability | $17.90 \%$ | $17.28 \%$ | $18.38 \%$ |
| Wages | 1.40 | 1.47 | 1.35 |
|  | $(0.80)$ | $(0.97)$ | $(0.64)$ |
| Move Firms within 2-Digit Industry |  |  |  |
| Probability conditional on remaining in data | $10.54 \%$ | $7.45 \%$ | $12.79 \%$ |
| Unconditional Probability | $2.52 \%$ | $1.73 \%$ | $3.13 \%$ |
| Wages | 1.41 | 1.58 | 1.33 |
|  | $(0.72)$ | $(0.90)$ | $(0.62)$ |

Move to Different Industry (High-Skill Sector)
Probability conditional on remaining in data
Unconditional Probability

| $7.53 \%$ | $7.82 \%$ | $7.32 \%$ |
| :---: | :---: | :---: |
| $1.80 \%$ | $1.82 \%$ | $1.79 \%$ |
| 1.50 | 1.51 | 1.44 |
| $(1.01)$ | $(1.42)$ | $(0.63)$ |


| Move to Different Industry (Low-Skill Sector) |  |  |  |
| :--- | :---: | :---: | :---: |
| Probability conditional on remaining in data | $7.13 \%$ | $10.34 \%$ | $4.79 \%$ |
| Unconditional Probability | $1.71 \%$ | $2.40 \%$ | $1.17 \%$ |
| Wages | 1.37 | 1.45 | 1.33 |
|  | $(0.67)$ | $(0.75)$ | $(0.68)$ |
| Observations | 21,376 | 9017 | 12,359 |

[^2]TABLE BV
Occupational distribution by sector.

|  | $(1)$ |  |
| :--- | :---: | :---: |
| Fraction of Workers in |  |  |
|  | Low-Skill Sectors |  |
| Professionals (0,1) | 0.023 | High-Skill Sectors |
| Legislative Bodies/Directors (2) | 0.041 | 0.050 |
| Office Workers (3) | 0.057 | 0.056 |
| Sales/Trade Workers (4) | 0.102 | 0.087 |
| Hospitality/Service Workers, etc. (5) | 0.065 | 0.087 |
| Agricultural Workers (6) | 0.026 | 0.025 |
| Mining/Industrial Workers (7) | 0.458 | 0.005 |
| Electricians/Mecchanics, etc. (8) | 0.128 | 0.062 |
| Craftsmen/Construction Workers, etc. (9) | 0.101 | 0.365 |
| Observations | 109,284 | 0.262 |

Note: The table shows the fraction of workers in different occupations by sector, based on data from the Colombian Household Survey (ECH). Occupations are categorized by 1-digit occupation codes (shown in parantheses).


Figure B1.-Wages Before and After Completing the Apprenticeship. Notes: The figure shows wages of apprentices who graduate during the sample period versus a control group of apprentices who do not graduate during this time, based on the PILA data. Time is measured in months relative to time of completing the apprenticeship. Wages are shown in units relative to minimum wage.


Figure B2.-Wages Before and After Completing the Apprenticeship by Sector. Notes: The figure shows the mean and standard deviation of the wages of apprentices who graduate during the sample period versus a control group of apprentices who do not graduate during this time, based on the PILA data. Time is measured in months relative to time of completing the apprenticeship. Wages are shown in units relative to minimum wage.

## APPENDIX C: Reduced-Form Results With Alternative Sector Classification

In this Appendix, we replicate the main reduced-form results from Section 3 with an alternative sector classification. We focus only on those sectors which can be unambiguously classified as high- or low-skill according to the various skill proxies shown in Appendix Table AI. The set of "clearly high-skill" sectors includes paper/editorial, metallic products, machinery, and chemical products, which are classified as high-skill according to all or virtually all proxies. The set of "clearly low-skill" sectors includes wood products, textile, and food/beverage, all of which are classified as low-skill according to all or virtually all proxies. Thus, compared to the main reduced-form results, two more ambiguous sectors are excluded, namely other manufacturing and mineral nonmetallic products.


Figure C1.-Bunching Responses with Alternative Sector Definition. Notes: The figure replicates main text Figure 3, excluding firms in the other manufacturing and mineral non-metallic sectors, for which the assignment to high- and low-skill is less clear. The figure shows the distribution of the number of full-time workers in clearly high-skill and clearly low-skill sectors post-reform (2003-2009), using a bin size of one. The dashed vertical lines denote the regulation thresholds. The solid thin line shows the fitted counterfactual density. Excess mass $b$ and missing mass $m$ are reported at each threshold, with bootstrapped standard errors in parentheses.

(a) Fact 2-Number of Apprentices by Firm Size


Firm Size (Number of Full-Time Workers)

- High-skill $\diamond$ Low-skill
(b) Fact 3-Share of Firms Paying Fees by Firm Size

Figure C2.-Number of Apprentices and Share of Firms Paying Fees with Alternative Sector Definitions. Notes: The figure replicates main text Figure 4, excluding firms in the other manufacturing and mineral nonmetallic sectors, for which the assignment to high- and low-skill is less clear. Panel (a) shows the average number of apprentices by firm size in clearly high-skill and clearly low-skill sector firms. The horizontal dashed lines show the minimum and maximum apprentice quotas, and the vertical dashed lines denote the regulation thresholds. Panel (b) shows the fraction of firms paying fees by firm size in clearly high-skill and clearly low-skill sectors. Both panels pool the post-reform years 2003 to 2009.

## APPENDIX D: Model Proofs and Extensions

## D.1. Equilibrium and Proofs

Here, we characterize the equilibrium with and without regulation, show some additional theoretical results and provide a proof of Proposition 1.

Assumption 1 on $f$ allows us to further characterize the optimal number of workers and apprentices, guaranteeing the existence and uniqueness of the solution. Conditions (i) and (ii) ensure the existence of a unique solution with $n>0$. Condition (iii) implies that the optimal number of workers $n^{*}\left(z, t_{a}\right)$ and apprentices $n_{a}^{*}\left(z, t_{a}\right)$ are nondecreasing in $z$. In other words, firms with higher managerial ability $z$ are larger. We formalize these claims in Lemma 1.

Lemma 1: Assumption 1 implies unique labor demands $n^{*}\left(z, t_{a}\right)>0$ and $n_{a}^{*}\left(z, t_{a}\right) \geq 0$ solving firm z's optimization problem (1). Moreover, these labor demands are nondecreasing in the managerial ability, $\frac{\partial n^{*}}{\partial z} \geq 0$ and $\frac{\partial n_{a}^{*}}{\partial z} \geq 0$.

Proof: Similar to standard production theory, homogeneity of degree $\gamma \in(0,1)$ in $\left(l, l_{a}\right)$ implies concavity (and hence quasi-concavity) of the production function and the existence of the solution $l^{*}(z), l_{a}^{*}(z)$. Additionally, the Inada condition on $n$ guarantees that the solution is unique. From these labor demands, we compute the optimal number of workers $n^{*}$ and apprentices $n_{a}^{*}, n^{*}=l^{*}+t_{a} n_{a}^{*}$, and $n_{a}^{*}=\frac{l_{a}^{*}}{\zeta_{a}}$.

Now, since the cross-derivatives are nonnegative, monotone comparative statics imply $\frac{\partial l^{*}(z)}{\partial z} \geq 0$ and $\frac{\partial l_{a}^{*}(z)}{\partial z} \geq 0$. Thus, $\frac{\partial n^{*}}{\partial z}, \frac{\partial n_{a}^{*}}{\partial z} \geq 0$.
Q.E.D.

At an interior solution, the marginal rate of substitution between the two types of labor is equal to the ratio of marginal labor costs

$$
-\frac{\frac{\partial f}{\partial l}}{\frac{\partial f}{\partial l_{a}}}=-\frac{w \zeta_{a}}{w_{a}+t_{a} w}
$$

We can use the FOCs to analyze how wages or required training time affect the optimal labor allocation decision. As usual, an increase in the relative wage of apprentices lowers demand for apprentices. Similarly, an increase in training costs decreases the demand for apprentices.

Lemma 2: Suppose that Assumption 1 holds. Then $\frac{n_{a}}{n}$ is weakly decreasing in $w_{a}$ and $t_{a}$, and weakly increasing in $w$.

PROOF: At an interior solution, from the firm's optimization problem,

$$
\frac{d l_{a}}{d l}=-\frac{\frac{\partial f}{\partial l}}{\frac{\partial f}{\partial l_{a}}}=-\frac{w \zeta_{a}}{w_{a}+t_{a} w}
$$

Let $W=\frac{w_{\zeta a}}{w_{a}+t_{a} w}$ denote the ratio of the price of workers' and apprentices' labor. In equilibrium, if $W$ increases, then $\frac{d l_{a}}{d l}$ decreases. Since $f$ is homogeneous of degree $\gamma \in$ $(0,1)$ in $\left(l, l_{a}\right)$, this means $l_{a}^{*} / l^{*}$ increases.

Now, $\frac{n_{a}^{*}}{n^{*}}=\frac{1}{\zeta_{a} \bar{L}_{a}^{*}+t_{a}}$. All comparative static results follow from this equation and the previous observations.

If $w_{a}$ increases, then $W$ decreases, so $l_{a}^{*} / l$ decreases, implying $n_{a}^{*} / n^{*}$ also decreases. Similarly, an increase in $t_{a}$ implies $W$ and $l_{a}^{*} / l^{*}$ decrease. Now, this increase in $t_{a}$ decreases $n_{a}^{*} / n^{*}$ directly and indirectly through $l_{a}^{*} / l^{*}$, so $n_{a}^{*} / n^{*}$ decreases. At the boundary, $n^{*}=t_{a} n_{a}^{*}$, and $\frac{n_{a}^{*}}{n^{*}}=\frac{1}{t_{a}}$ is also decreasing in $t_{a}$. The analogous logic applies to $w$. If $w$ increases, $W$ and $l_{a}^{*} / l^{*}$ also rise, implying $n_{a}^{*} / n^{*}$ increases.
Q.E.D.

## Equilibrium

All individuals in the economy are infinitely lived, have a common utility function, and are endowed with a unit of labor which they supply inelastically. Each individual $i$ maximizes their lifetime utility

$$
\begin{equation*}
\max _{\left(c_{t^{\prime}}\right.} \sum_{t=1}^{\infty} \beta^{t} u\left(c_{t}\right) \quad \text { s.t. } \sum_{k=1}^{K} p_{t}^{k} c_{t}^{k}=I_{t}^{i} \forall t, \tag{6}
\end{equation*}
$$

where $I_{i, t}$ denotes income in period $t$. Note that as for firms', individuals' decisions are static. Solving this problem implies the usual optimality conditions, $\frac{\partial u_{t} / \partial c_{t}^{j}}{\partial u_{t} / \partial c_{t}^{k}}=\frac{p_{t}^{k}}{p_{t}^{j}}$ and $\sum_{k=1}^{K} p_{t}^{k} c_{t}^{k}=I_{t}^{i} \forall k, t$. Assuming $u(\cdot)$ is quasiconcave, let $c_{t}^{*}\left(I_{t}^{i} ; p_{t}\right)=\left(c_{t}^{1 *}\left(I_{t}^{i} ; p_{t}\right), \ldots\right.$, $c_{t}^{K *}\left(I_{t}^{i} ; p_{t}\right)$ ) be the solution to individual $i$ 's optimization problem in period $t$. We use the goods market clearing conditions for each sector $k$ to determine prices $p_{t}=\left(p_{t}^{1}, \ldots, p_{t}^{K}\right)$,

$$
C_{t}^{k}\left(p_{t}\right)+C_{a, t}^{k}\left(p_{t}\right)+C_{f, t}^{k}\left(p_{t}\right)=Y_{t}^{k}\left(p_{t}\right) \quad \forall k
$$

where $C_{t}^{k}\left(p_{t}\right)$ is workers' aggregate demand for good $k, C_{a, t}^{k}$ is apprentices' aggregate demand, $C_{f, t}^{k}$ is firm owners' aggregate demand and $Y_{t}^{k}\left(p_{t}\right)$ is aggregate production of good $k$ :

$$
\begin{aligned}
C_{t}^{k}\left(p_{t}\right) & :=\sum_{j=1}^{K} L_{t}^{j} c_{t}^{k *}\left(w_{t}^{j} ; p_{t}\right), \quad C_{a, t}^{k}:=\sum_{j=1}^{K} L_{a, t}^{j} c_{t}^{k *}\left(w_{a, t} ; p_{t}\right), \\
C_{f, t}^{k} & :=\sum_{j=1}^{K} F_{t}^{j} \iint c_{t}^{k *}\left(\pi_{t}^{j}\left(z, t_{a}\right) ; p_{t}\right) d \mathcal{Z}(z) d \mathcal{T}\left(t_{a}\right), \\
Y_{t}^{k}\left(p_{t}\right) & :=F_{t}^{k} \iint y_{t}^{k *}\left(p_{t} ; z, t_{a}\right) d \mathcal{Z}(z) d \mathcal{T}\left(t_{a}\right),
\end{aligned}
$$

where $F_{t}^{j}$ denotes the number of firms in sector $j$ at time $t$.
DEFINITION 1: A competitive equilibrium is given by wages $\left(\left(w_{t}^{k *}\right)_{k}, w_{a, t}^{k *}\right)_{t}$ and prices $p_{t}^{k}$ for each sector $k$ and each period $t$; and quantities of unemployed workers and untrained apprentices $\left(\left(U_{t}^{k *}\right)_{k}, U_{a, t}^{*}\right)_{t}$, labor demands $\left(n_{t}^{k *}\left(z, t_{a}\right), n_{a, t}^{k *}\left(z, t_{a}\right)\right)$ for each firm $\left(z, t_{a}\right)$ and consumption $c_{t}^{k *}$ such that
(i) firms solve the optimization problem (1),
(ii) wage restrictions are satisfied, $w_{t}^{k *} \geq w_{a, t}^{k *} \geq w_{\min }$, and labor markets clear with $U_{t}^{k *} \geq 0$ and $U_{a, t}^{*} \geq 0 \forall t, k$.
(iii) apprentices increase total labor in each period and all sectors, as in (2),
(iv) individuals maximize utility (6) and,
(iv) the goods market clear for each sector.

Note that unemployment and untrained apprentices exist whenever the wage restrictions are binding.

## Equilibrium With Regulation

Lemma 3 characterizes the solution to the optimization problem with regulation relative to the problem without regulation. We use this characterization for the proof of Proposition 1.

Let $n^{*}\left(z, t_{a}\right), n_{a}^{*}\left(z, t_{a}\right)$ denote the optimal number of workers and apprentices a firm with managerial ability $z$ and training costs $t_{a}$ hires when solving the maximization problem (1) (without regulation); and $n^{r}\left(z, t_{a}\right)$ and $n_{a}^{r}\left(z, t_{a}\right)$ denote the optimal number of workers and apprentices the firm hires when solving (3) (with regulation).

LEMMA 3: Let $j \geq 1$ be such that $n^{*}\left(z, t_{a}\right) \in\left[N_{j-1}, N_{j}\right)$ and $\left(\underline{n}_{a}^{j}, \bar{n}_{a}^{j}\right)$ be the corresponding minimum and maximum quotas.
i. If $n_{a}^{*}\left(z, t_{a}\right)>\bar{n}_{a}^{j}$, then

- $\exists k \geq j$ such that $n^{r}\left(z, t_{a}\right)=N_{k}$ and $n_{a}^{r}\left(z, t_{a}\right)>\bar{n}_{a}^{j}$ (increase size to get more apprentices), or
- $n_{a}^{r}\left(z, t_{a}\right)=\bar{n}_{a}^{j}$ and $n^{r}\left(z, t_{a}\right)<n^{*}\left(z, t_{a}\right)$ (bounded by maximum quota).
ii. If $n_{a}^{*}\left(z, t_{a}\right) \in\left[\underline{n}_{a}^{j}, \bar{n}_{a}^{j}\right]$, then $n^{r}\left(z, t_{a}\right)=n^{*}\left(z, t_{a}\right)$ and $n_{a}^{r}\left(z, t_{a}\right)=n_{a}^{*}\left(z, t_{a}\right)$.
iii. If $n_{a}^{*}\left(z, t_{a}\right)<\underline{n}_{a}^{j}$, then
- $\exists k<j$ such that $n^{r}\left(z, t_{a}\right)=N_{k}-\varepsilon($ with $\varepsilon \rightarrow 0)$ and $n_{a}^{r}\left(z, t_{a}\right)<\bar{n}_{a}^{j}$ (reduce size to avoid apprentices),
- $n^{r}\left(z, t_{a}\right) \geq n^{*}\left(z, t_{a}\right)$ and $n_{a}^{r}\left(z, t_{a}\right)<\underline{n}_{a}^{j}$ and $d_{f_{a}}=1$ (pay the fee to avoid apprentices) or
- $n_{a}^{r}\left(z, t_{a}\right)=\underline{n}_{a}^{j}$ (bounded by the minimum quota).

Proof: Pick any firm $z>0$ and $t_{a} \geq 0$. Denote by $\pi(N)$ the maximum profit function when the number of workers is fixed to $N$ and $\pi\left(N_{a}\right)$ when the number of apprentices is fixed to $N_{a}$ :

$$
\begin{aligned}
\pi(N) & =\max _{l_{a} \geq 0} p f\left(N-\frac{t_{a}}{\zeta_{a}} l_{a}, l_{a}\right)-w N-\frac{w_{a}+t_{a} w}{\zeta_{a}} l_{a}, \\
\pi\left(N_{a}\right) & =\max _{l \geq 0} p f\left(l, N_{a}\right)-w l-\left(w_{a}+t_{a} w\right) N_{a}
\end{aligned}
$$

To simplify notation, define $\tilde{w}_{a}=\frac{w_{a}+t_{a} w}{\zeta_{a}}$ and $\tilde{t}_{a}=t_{a} / \zeta_{a}$. We also use the subindex notation of partial derivatives to economize on writing, $f_{x}:=\frac{\partial f}{\partial x}$.

First, we show that $\pi(N)$ and $\pi\left(N_{a}\right)$ are concave. Let us start with $\pi\left(N_{a}\right)$. Using the envelope theorem, $\frac{\partial \pi\left(N_{a}\right)}{\partial N_{a}}=p f_{l_{a}} \zeta_{a}-\left(w_{a}+t_{a} w\right)$. We can differentiate this expression again with respect to $N_{a}$ to obtain

$$
\begin{equation*}
\frac{\partial^{2} \pi\left(N_{a}\right)}{\partial N_{a}^{2}}=p\left(f_{l_{a} l} \frac{d l^{r}}{d N_{a}}+f_{l_{a} l_{a}} \zeta_{a}\right) \tag{7}
\end{equation*}
$$

where $l^{r}$ solves the FOC of the optimization problem with fixed $N_{a}, p f_{l}\left(l^{r}, N_{a}\right)=w$. Assumption 1 implies the existence and uniqueness of this solution $l^{r}$. Totally differentiating the FOC with respect to $l^{r}$ and $N_{a}$ implies $\frac{d l^{r}}{d N_{a}}=-\frac{f_{l_{a}}}{f_{l l}} \zeta_{a} \geq 0$, as $f_{l l_{a}} \geq 0$ and $f_{l l} \leq 0$. Replacing this derivative in (7) yields

$$
\frac{\partial^{2} \pi\left(N_{a}\right)}{\partial N_{a}^{2}}=p\left(f_{l_{a} l}\left(-\frac{f_{l_{a}}}{f_{l l}} \zeta_{a}\right)+f_{l_{a} l_{a}} \zeta_{a}\right)=\frac{p \zeta_{a}}{f_{l l}}\left(f_{l l} f_{l a l_{a}}-f_{l l_{a}}^{2}\right) \leq 0
$$

since concavity of $f$ in $l, l_{a}$ implies $f_{l_{a} l_{a}} \leq 0$ and $\left(f_{l l} f_{l_{a} l_{a}}-f_{l l_{a}}^{2}\right) \geq 0$. Thus, $\pi\left(N_{a}\right)$ is concave in $N_{a}$.

Importantly, this function is maximized at $\left(n^{*}, n_{a}^{*}\right)$. So if we choose $n_{a}$ away from $n_{a}^{*}$, profits will decrease. In Case (i), if $n_{a}^{*}>\bar{n}_{a}^{j}$, whenever the firm stays in the same $j$ th regulation bracket, it chooses the feasible number of apprentices that is closest to $n_{a}^{*}$. This means the upper bound is binding, $n_{a}^{r}=\bar{n}_{a}^{j}$. Moreover, since $\frac{d l^{r}}{d N_{a}} \geq 0$ we know $l^{r}<l^{*}$ (given $n_{a}^{*}>\bar{n}_{a}^{j}$ ). This implies $n^{r}=l^{r}+t_{a} n^{r}=l^{r}+t_{a} \bar{n}_{a}^{j}<n^{*}$.

Similarly, we can show that $\pi(N)$ is concave. In this case,

$$
\frac{\partial^{2} \pi(N)}{\partial N^{2}}=p f_{l l}\left(1-\tilde{t}_{a} \frac{d l_{a}^{r}}{d N}\right)+p f_{l_{l}} \frac{d l_{a}^{r}}{d N}
$$

Considering the FOC and totally differentiating,

$$
\frac{d l_{a}}{d N}=-\frac{f_{l_{a} l}-f_{l l} \tilde{t}_{a}}{f_{l l} \tilde{t}_{a}^{2}-2 f_{l l_{a}} \tilde{t}_{a}+f_{l_{a} l_{a}}} \geq 0
$$

Substituting in the previous equation,

$$
\frac{\partial^{2} \pi(N)}{\partial N^{2}}=p \frac{\left(f_{l l} f_{l_{a} l_{a}}-f_{l_{a}}^{2}\right)}{f_{l l} \tilde{t}_{a}^{2}-2 f_{l_{a}} \tilde{t}_{a}+f_{l_{a} l_{a}}} \leq 0
$$

The last inequality stems again from the concavity of $f$, as $f_{l l} f_{l a} l_{a}-f_{l l_{a}}^{2}>0, f_{l l}<0$, and $f_{l_{a} l_{a}}<0$.

This proves $\pi(N)$ is concave. Using a similar argument, the firm wants to get as close as possible to the optimal labor demands $\left(n^{*}, n_{a}^{*}\right)$. However, now we also have to compare subsequent thresholds $N_{k}$ for $k \geq j+1$, as $n_{a}^{r}\left(z, t_{a}\right)$ might still be larger than $\bar{n}_{a}^{j+1}$, so a firm might want to jump multiple thresholds to get a higher number of apprentices. In all of these cases, the firm chooses the number of workers at the threshold $N_{k}$ as it is the closest to the optimal number of workers that allows the firm to get $n_{a}^{r} \in\left[\bar{n}_{a}^{k-1}, \bar{n}_{a}^{k}\right]$. The optimal number of apprentices in this case is $n_{a}^{r}\left(z, t_{a}\right)>\bar{n}_{a}^{j}$.

Case (ii) is immediate as the unconstrained optimum is within the regulation bounds, so the firm does not change its optimal decision.

For Case (iii), the proof is analogous to Case (i) whenever firms bunch just below a threshold or choose apprentices at the minimum quota. It remains to show that for relatively low $\phi_{a}>0$, some firms prefer to pay the fee instead of hiring the minimum number of required apprentices.

To see this, suppose $n_{a}^{*}<\underline{n}_{a}^{j}$ and define $\pi^{*}\left(\phi_{a}\right):=p f\left(l^{*}, l_{a}^{*}\right)-w l^{*}-\tilde{w}_{a} l_{a}^{*}-\phi_{a}\left(\underline{n}_{a}^{j}-n_{a}^{*}\right)$, the profits when hiring labor as optimal without regulation, where $\tilde{w}_{a}:=\frac{w_{a}+t_{a} w}{\zeta_{a}}$. Note that
the optimal choice of workers and apprentices when paying the fee has to yield larger or equal profits,

$$
\pi^{j}\left(\phi_{a}\right):=\max _{l_{a}, l \geq 0} p f\left(l, l_{a}\right)-w l-\tilde{w}_{a} l_{a}-\phi_{a}\left(\underline{n}_{a}^{j}-\frac{l_{a}}{\zeta_{a}}\right) \geq \pi^{*}\left(\phi_{a}\right)
$$

Now, we know $\pi^{*}:=p f\left(l^{*}, l_{a}^{*}\right)-w l^{*}-\tilde{w}_{a} l_{a}^{*} \geq \pi(N)$ and $\pi^{*} \geq \pi\left(N_{a}\right)$, for any $N, N_{a} \geq 0$. Also, $\pi^{j}\left(\phi_{a}\right)$ is continuous in $\phi_{a}$, and $\lim _{\phi_{a} \rightarrow 0} \pi^{*}\left(\phi_{a}\right)=\pi^{*}$. Hence, there exists a small enough $\tilde{\phi}_{a}>0$, such that $\pi^{j}\left(\tilde{\phi}_{a}\right) \geq \pi(N)$ and $\pi^{j}\left(\tilde{\phi}_{a}\right) \geq \pi\left(N_{a}\right)$.
Q.E.D.

Lemma 4: Suppose Assumption 1 holds, except $f$ does not necessarily have constant returns to scale $(C R S)$ in $\left(l, l_{a}, z\right)$. For each firm $z$, there exists $A(z)>0$ such that $l_{a}^{*}=A(z) l^{*}$ :
i. If $A^{\prime}(z)>0$, the parametric mapping $\left(n^{*}(z), n_{a}^{*}(z)\right)$ is strictly convex.
ii. If $A^{\prime}(z)=0$, the parametric mapping $\left(n^{*}(z), n_{a}^{*}(z)\right)$ is linear.
iii. If $A^{\prime}(z)<0$, the parametric mapping $\left(n^{*}(z), n_{a}^{*}(z)\right)$ is strictly concave.

PROOF: Take any firm $z>0$. First, let us show that $l_{a}^{*}=A(z) l^{*}$. Since $f$ is homogenous of degree $\gamma$, then $\frac{\partial f}{\partial l}$ and $\frac{\partial f}{\partial l_{a}}$ are homogenous of degree $\gamma-1$. Thus, for any constant $k>0$,

$$
\frac{\frac{\partial f}{\partial l}\left(k l, k l_{a} ; z\right)}{\frac{\partial f}{\partial l_{a}}\left(k l, k l_{a} ; z\right)}=\frac{k^{\gamma-1} \frac{\partial f}{\partial l}\left(l, l_{a} ; z\right)}{k^{\gamma-1} \frac{\partial f}{\partial l_{a}}\left(l, l_{a} ; z\right)}=\frac{\frac{\partial f}{\partial l}\left(l, l_{a} ; z\right)}{\frac{\partial f}{\partial l_{a}}\left(l, l_{a} ; z\right)}
$$

So the derivatives of the isoquants are constant along any ray starting from the origin. Since $\gamma \in(0,1)$ implies the production function is quasiconcave and the Inada condition holds for workers, there is only one point $\left(l^{*}, l_{a}^{*}\right)$ such that $-\frac{\frac{\partial f}{\frac{d}{l}\left(l^{*}, l_{a}^{*}\right)}}{\frac{\partial f}{\partial l_{a}}\left(l^{*}, l_{a}^{*}\right)}=-\frac{w \zeta_{a}}{w_{a}+t_{a} w}$. Together this implies $l_{a} / l$ is constant whenever the derivative of the isoquant is constant. Hence, $\frac{l_{a}^{*}}{l^{*}}=A(z)$ for some $A(z)>0$.

Now note that since $l=n-t_{a} n_{a}$ and $l_{a}=\zeta_{a} n_{a}, \frac{n_{a}^{*}}{n^{*}}=\frac{1}{\zeta_{a} A(z)^{-1}+t_{a}}$. Call this last term $B(z)$. Hence, $A^{\prime}(z)>0 \Longleftrightarrow B^{\prime}(z)>0$.

From the equation above, $\frac{d n_{a}^{*}}{d n^{*}}=B(z), \forall z$. Thus, if $B(z)$ is increasing in $z$, then $\frac{d n_{a}^{*}}{d n^{*}}$ is increasing in $z$. From Lemma (1), $\frac{d n^{*}}{d z}>0$, which implies that the parametric mapping $n_{a}^{*}\left(n^{*}\right)$ is convex. Similarly, if $A^{\prime}(z)=0 \Rightarrow B^{\prime}(z)=0$ and so the derivative is constant for any $z, \frac{d n_{a}^{*}}{d n^{*}}=B \in \mathbb{R}_{+}, \forall z$. This means the parametric mapping is linear. Finally, if $A^{\prime}(z)<$ $0 \Rightarrow B^{\prime}(z)<0$, then $\frac{d n_{a}^{*}}{d n^{*}}$ is decreasing in $z$ and hence $n_{a}^{*}\left(n^{*}\right)$ is concave.
Q.E.D.

Corollary 1: Under Assumption $1, n_{a}^{*}=B n^{*}$, where $B$ does not depend on $z$.
Proof: Assuming CRS of $f$ on $\left(l, l_{a}, z\right)$ implies $\frac{(\partial f / \partial l)}{\left(\partial f / \partial l_{a}\right)}$ is homogenous of degree 0 in $z$. So Case (ii) of Lemma 4 applies $l_{a}^{*}=A l^{*}$ for some $A \geq 0$ independent of $z$. Using the same argument in Lemma 4, $n_{a}^{*}=B n^{*}$ for $B=\frac{1}{\zeta_{a} A^{-1}+t_{a}}$.
Q.E.D.

Proposition 1: Suppose Assumption 1 holds and firms solve the maximization problem with regulation (3). Then $\forall z \geq 0$,

Case 1: there exist $\left(\frac{\overline{w_{a}}}{w}, \bar{t}_{a}\right)$ such that for $\frac{w_{a}}{w} \leq \frac{\overline{w_{a}}}{w}$ and $t_{a} \leq \bar{t}_{a}$,
i. the number of apprentices without regulation is $n_{a}^{*}=B_{u} n^{*}$ and is above the maximum quota, $n_{a}^{*}\left(z, t_{a}\right)>\bar{n}_{a}^{j}$.
ii. there exist cutoffs $\left\{z_{b}^{j}, z_{r}^{j}\right\}$ such that firms $z \in\left[z_{b}^{j}, z_{r}^{j}\right]$ increase their size to threshold $N_{k}$ with $k \geq j$.
iii. firms choose maximum number of apprentices $n_{a}^{r}=\bar{n}_{a}^{j}$.
iv. firms never pay the fee.

Case 2: there exist $\left(\frac{w_{a}}{w}, \underline{t}_{a}\right)$ such that for $\frac{w_{a}}{w} \geq \frac{w_{a}}{w}$ or $t_{a} \geq \underline{t}_{a}$,
i. the number $\overline{\text { of }}$ apprentices without regulation is $n_{a}^{*}=B_{s} n^{*}$ and is below the minimum quota, $n_{a}^{*}\left(z, t_{a}\right)<\underline{n}_{a}^{j}$.
ii. there exist cutoffs $\left\{z_{b}^{j}, z_{r}^{j}\right\}$ such that firms $z \in\left[z_{b}^{j}, z_{r}^{j}\right]$ reduce their size $\epsilon$ below threshold $N_{k}$ with $k<j$.
iii. firms that increase the number of apprentices choose the minimum number $\underline{n}_{a}^{j}$.
iv. there exists $\bar{\phi}_{a}>0$ such that for $\phi_{a} \leq \bar{\phi}_{a}$, there is an additional cutoffs $z_{f}^{j}$ where firms $z \in\left(z_{r}^{j}, z_{f}^{j}\right]$ choose to pay the fee.

Proof: Let $z \geq 0$. Assumption 1 and Corollary 1 imply that the solution without regulation is $n_{a}^{*}=B n^{*}$ for some $B \geq 0$ that does not depend on $z$. Lemma 2 implies that $B$ is a continuous nonincreasing function of $\frac{w_{a}}{w}$ and $t_{a}$.

Let us start with Case 1 . When $\frac{w_{a}}{w}$ and $t_{a}$ approach 0 , the optimal relative number of apprentices $n_{a}^{*} / n^{*}$ is unbounded. This means below some threshold $\left(\frac{\overline{w_{a}}}{w}, \bar{t}_{a}\right), n_{a}^{*}\left(z, t_{a}\right)>\bar{n}_{a}^{j}$ with $n^{*}\left(z, t_{a}\right) \in\left[N_{j-1}, N_{j}\right)$ (part i). Now, for any $t_{a}<\bar{t}_{a}, \frac{w_{a}}{w} \leq \frac{\overline{w_{a}}}{w}$, Lemma 1 implies $n_{a}^{*}$ and $n^{*}$ are increasing and continuous in $z$. After some threshold $z_{b}^{j}$, by Lemma 3, $\left(N_{k}, n_{a}\left(N_{k}\right)\right)$ for some $k \geq j$ is closer to the optimal labor input ( $n_{a}^{*}, n^{*}$ ), and hence firm ( $z, t_{a}$ ) with $z \geq z_{b}^{j}$ will bunch at this threshold $N_{k}$. Let $z_{r}^{j}$ be productivity such that $N_{k}$ is the optimal number of workers without regulation for firm $\left(z_{r}^{j}, t_{a}\right)$. Then firms beyond $z>z_{r}^{j}$ do not increase their size relative to the equilibrium with no regulation, completing the proof of part (ii). Since $n_{a}^{*}>\bar{n}_{a}^{j}$ by Lemma 3, firms choose $n_{a}^{r}=\bar{n}_{a}^{j}$ (part iii). And finally, firms never pay the fee as the number of desired apprentices is above the maximum quota (part iv).

The proof of Case 2 is analogous. When either $\frac{w_{a}}{w}$ or $t_{a}$ tend to $\infty$, the optimal relative number of apprentices $n_{a}^{*} / n^{*}$ converges to zero. Note that for $\phi \rightarrow 0$, we can use the same argument as in the proof of Lemma 3 in order to prove that firms prefer to pay the fee instead of hiring the minimum quota of apprentices.
Q.E.D.

## Negative Marginal Productivity of Apprentices

In the following, we show that the production function has to allow for apprentices having negative marginal productivity in firms that choose to pay the fee. Consider a standard production function $\tilde{f}\left(n, n_{a} ; z\right)$ combining managerial ability $z$ with labor input from workers and apprentices. Let us compare two scenarios based on the components of the regulation. First, suppose that firms are required to train at least $\underline{n}_{a}$ apprentices paying a wage of $w_{a}^{\min }$. Alternatively, firms can pay a fee $\phi_{a}$ per required apprentice.

When a firm $z$ chooses to train the apprentices, it solves

$$
\pi_{a}(z):=\max _{n, n_{a}} p \tilde{f}\left(n, n_{a} ; z\right)-w n-w_{a}^{\min } n_{a}, \quad n_{a} \geq \underline{n}_{a} .
$$

Instead, the firm could pay the fee and solve

$$
\pi_{f}(z):=\max _{n} p \tilde{f}(n, 0 ; z)-w n-\phi_{a} \underline{n}_{a} .
$$

In Proposition 2, we show that if $\phi_{a}>w_{a}^{\min }$, then firms choose to pay the fee only if the marginal productivity of apprentices is negative.

PROPOSITION 2: If $\phi_{a}>w_{a}^{\min }$, then $\pi_{f}(z)>\pi_{a}(z) \Rightarrow \frac{\partial \tilde{f}}{\partial n_{a}}<0$.
PROOF: Let $z \geq 0$ and $n_{0}^{*}$ denote the number of workers maximizing $\pi_{f}(z)$. By way of contradiction, suppose $\frac{\partial \tilde{f}}{\partial n_{a}} \geq 0$. In this case,

$$
\begin{aligned}
\pi_{a}(z) & \geq p \tilde{f}\left(n_{0}^{*}, \underline{n}_{a} ; z\right)-w n_{0}^{*}-w_{a} \underline{n}_{a} \geq p \tilde{f}\left(n_{0}^{*}, 0 ; z\right)-w n_{0}^{*}-w_{a} \underline{n}_{a} \\
& >p \tilde{f}\left(n_{0}^{*}, 0 ; z\right)-w n_{0}^{*}-\phi_{a} \underline{n}_{a}=\pi_{f}(z)
\end{aligned}
$$

where the first two inequalities are implied by $\pi_{a}(z)$ being the maximum profit function and $\frac{\partial \tilde{f}}{\partial n_{a}} \geq 0$, and the last inequality is due to $\phi_{a}>w_{a}$. This contradicts $\pi_{f}(z)>\pi_{a}(z)$. Therefore, $\pi_{f}>\pi_{a} \Rightarrow \frac{\partial \tilde{f}}{\partial n_{a}}<0$. Q.E.D.

In other words, if apprentices have positive productivity and it is cheaper to hire an apprentice than paying the fee, firms choose to hire the apprentice. In the model, we allow for negative marginal revenue product of apprentices simply by adding training costs.

## D.2. Additional Inputs

In this section, we describe an extension of the model adding other inputs. We discuss a simple example to illustrate the results based on the baseline model.

Suppose there is an additional input $x$, with price $w_{x}$, that firms choose in each period. First, we consider the firm problem without regulation. Consider a simple Cobb-Douglas specification

$$
\max _{n, n_{a}, x} p z^{1-\gamma}\left(n-t_{a} n_{a}+\zeta_{a} n_{a}\right)^{\gamma_{l}} x^{\gamma_{x}}-w n-w_{a} n_{a}-w_{x} x \quad \text { s.t. } t_{a} n_{a} \leq n
$$

where $\gamma_{l}$ is the output elasticity of labor, and $\gamma_{x}$ the output elasticity of input $x$. Suppose the production function has constant returns to scale on $\left(z, n, n_{a}, x\right)$, and thus $\gamma_{l}+\gamma_{x}=\gamma$. As in the baseline model, linearity in labor input implies that there are corner solutions.

A firm $\left(z, t_{a}\right)$ avoids apprentices whenever $w<\frac{t_{a} w+w_{a}}{\zeta_{a}}$. In that case, from the FOC, $x=\frac{w}{w_{x}} \frac{\gamma_{x}}{\gamma_{l}} n=: \mathcal{X} n$. So the optimal input demands are $n_{a}^{*}=0, n^{*}=\left(\frac{p \gamma_{l} \mathcal{X}^{\gamma_{x}}}{w}\right)^{\frac{1}{1-\gamma}} z, x^{*}=\mathcal{X} n^{*}$. The corresponding output and profits are

$$
y^{*}=\left(\frac{p \gamma_{l}}{w}\right)^{\gamma /(1-\gamma)} \mathcal{X}^{\frac{\gamma_{x}}{1-\gamma}} z, \quad \pi^{*}=p^{\frac{1}{1-\gamma}}\left(\frac{\gamma_{l}}{w}\right)^{\gamma /(1-\gamma)} \mathcal{X}^{\frac{\gamma_{x}}{1-\gamma}}(1-\gamma) z
$$

On the other hand, if $w>\frac{t_{a} w+w_{a}}{\zeta_{a}}$ the firm seeks apprentices, so $x=\frac{w_{a}+t_{a} w}{w_{x}} \frac{\gamma_{x}}{\gamma_{l}} n_{a}=: \mathcal{X}_{a} n_{a}$. The optimal input demands are $n_{a}^{*}=\left(\frac{p \gamma_{l_{5}^{\prime}}^{\gamma_{a}^{\prime}} \mathcal{X}_{a}^{\gamma_{x}^{x}}}{w_{a}+t_{a} w}\right)^{\frac{1}{1-\gamma}} z, n^{*}=t_{a} n_{a}^{*}, x^{*}=\mathcal{X}_{a} n_{a}^{*}$, and output and profits are

$$
y^{*}=\left(\frac{p \gamma_{l}}{w_{a}+t_{a} w}\right)^{\gamma /(1-\gamma)} \zeta_{a}^{\frac{\gamma_{l}}{1-\gamma}} \mathcal{X}_{a}^{\frac{\gamma_{x}}{1-\gamma}} z, \quad \pi^{*}=p^{\frac{1}{1-\gamma}}\left(\frac{\gamma_{l}}{w_{a}+t_{a} w}\right)^{\gamma /(1-\gamma)} \zeta_{a}^{\frac{\gamma_{l}}{1-\gamma}} \mathcal{X}_{a}^{\frac{\gamma_{x}}{1-\gamma}}(1-\gamma) z
$$

Now, let us study the case of a particular threshold with regulation. Firms have the options of bunching at the threshold $N$, complying with the apprenticeship quotas by hiring the required number of apprentices $\underline{n}_{a}$, or paying the fee.

Suppose that a firm bunches at $N$ to avoid training:

$$
\begin{aligned}
& n^{r}=N, \quad n_{a}^{r}=0, \quad x^{r}=\left(\frac{\gamma_{x} N^{\gamma_{l}}}{w_{x}}\right)^{1 /\left(1-\gamma_{x}\right)} z^{\frac{1-\gamma}{1-\gamma_{x}}} \\
& y^{r}=\left(\frac{p \gamma_{x}}{w_{x}}\right)^{\frac{\gamma_{x}}{1-\gamma_{x}}} N^{\frac{\gamma-\gamma_{x}}{1-\gamma_{x}}} z^{\frac{1-\gamma}{1-\gamma_{x}}}, \quad \pi^{r}=p^{\frac{1}{1-\gamma_{x}}}\left(\frac{\gamma_{x}}{w_{x}}\right)^{\frac{\gamma_{x}}{1-\gamma_{x}}}\left(1-\gamma_{x}\right) N^{\frac{\gamma-\gamma_{x}}{1-\gamma_{x}}} z^{\frac{1-\gamma}{1-\gamma_{x}}},
\end{aligned}
$$

If the firm has to take $n_{a}$ apprentices instead, the same analysis applies:

$$
n^{r}=\left(\frac{p \gamma_{l} \mathcal{X}^{\gamma_{x}}}{w}\right)^{\frac{1}{1-\gamma}} z-\left(\zeta_{a}-t_{a}\right) n_{a}, \quad x^{r}=\mathcal{X} n^{r}
$$

Suppose the firm pays the fee:

$$
n_{a}^{r}=0, \quad n^{r}=\left(\frac{p \gamma_{l} \mathcal{X}^{\gamma_{x}}}{w}\right)^{\frac{1}{1-\gamma}} z, \quad x^{r}=\mathcal{X} n^{r}
$$

The corresponding output and profits are

$$
y^{r}=\left(\frac{p \gamma_{l}}{w}\right)^{\gamma /(1-\gamma)} \mathcal{X}^{\frac{\gamma_{x}}{1-\gamma}} z, \quad \pi^{r}=p^{\frac{1}{1-\gamma_{x}}}\left(\frac{\gamma_{l}}{w}\right)^{\gamma /(1-\gamma)} \mathcal{X}^{\frac{\gamma x}{1-\gamma}}(1-\gamma) z-\phi_{a} \underline{n}_{a} .
$$

From the equations above, we can see that the effect of adding other inputs is that firms have additional margins of substitution. Qualitatively, there are no differences to the baseline model used throughout the paper. However, this quantitatively affects the magnitude of firm responses. In terms of the estimation, the fit of the firm size distribution would be similar, but we would need information on the share of these other inputs in production to identify the parameters of the production function. Larger $\gamma_{x}$ implies firms would respond less to the regulation as the output elasticity with respect to labor decreases.

## D.3. Multiple Types of Workers

In this section, we describe an extension of the model including multiple types of workers. For clarity of exposition, let us suppose there are two types of workers, unskilled $u$ and skilled $s$. We characterize the equilibrium in the linear labor input case, combining these types of workers in a Cobb-Douglas function.

Suppose firms are characterized by managerial ability $z$ and training costs for each type of worker $t_{a}^{i}$ where $i \in\{u, s\}$. First, we study the case without regulation. A firm $\left(z, t_{a}^{u}, t_{a}^{s}\right)$ solves

$$
\max _{n^{i}, n_{a}^{i}} p z^{1-\gamma}\left(n^{u}+\left(\zeta_{a}^{u}-t_{a}^{u}\right) n_{a}^{u}\right)^{\gamma_{u}}\left(n^{s}+\left(\zeta_{a}^{s}-t_{a}^{s}\right) n_{a}^{s}\right)^{\gamma_{s}}-\sum_{i=u}^{s}\left(w^{i} n^{i}+w_{a} n_{a}^{i}\right) \quad \text { s.t. } t_{a}^{i} n_{a}^{i} \leq n^{i}, \forall i
$$

where $\sum_{i} \gamma_{i}=\gamma$, and $\zeta_{a}^{i} \in[0,1]$ denote the apprentices' productivity in each occupation.

As before, there are corner solutions. Firms avoid apprentices of type $i$ if $w^{i}<\frac{w^{i} t_{a}^{i}+w_{a}}{\zeta_{a}^{i}}$ : $n_{a}^{i}=0$ and $n^{i}=\left(\frac{p \gamma_{i} A_{j}}{w^{i}}\right)^{1 /(1-\gamma)} z$, where $A=\frac{w^{i}}{w^{j}} \frac{\gamma_{j}}{\gamma_{i}}$ and $i \neq j$. Firms seek apprentices of type $i$ if $w^{i}>\frac{w^{i} t_{a}^{i}+w_{a}}{\zeta_{a}^{i}}: n_{a}^{i}=\left(\frac{p \gamma_{i}\left(\zeta_{a}^{i} \gamma_{i} A^{\gamma_{j}}\right.}{w^{i} i_{a}^{i}+w_{a}}\right)^{1 /(1-\gamma)} z$ and $n^{i}=t_{a}^{i} n_{a}^{i}$.

Now let us consider the firm decision with regulation. Suppose the firm has to train $n_{a}$ apprentices due to quotas. First, we show that firms generically choose to train apprentices only in one occupation (by only one type of worker), depending on which one is relatively cheaper.

LEMMA 5: A firm chooses to train apprentices only in occupation $i^{*}=\arg \max _{i} w^{i}\left(\zeta_{a}^{i}-t_{a}^{i}\right)$.
PROOF: The net benefit of training an apprentice in occupation $i$ is $\zeta_{a}^{i} w^{i}-\left(w^{i} t_{a}^{i}+\right.$ $\left.w_{a}\right)=w^{i}\left(\zeta_{a}^{i}-t_{a}^{i}\right)-w_{a}$. Hence, firms choose to train apprentices only in occupation $i^{*}=$ $\arg \max _{i} w^{i}\left(\zeta_{a}^{i}-t_{a}^{i}\right)$.
Q.E.D.

Lemma 5 implies that we only have to compare the corner solutions to the choice of apprentices. Suppose the firm optimally chooses to train apprentices in occupation $i$. Let $x_{i}^{*}=p\left(\left(\frac{\gamma_{i}}{w^{i}}\right)^{1-\gamma_{j}}\left(\frac{\gamma_{j}}{w^{j}}\right)^{\gamma_{j}}\right)^{1 /(1-\gamma)} z$, then $n_{i}^{r}=x_{i}^{*}-\left(\zeta_{a}^{i}-t_{a}^{i}\right) n_{a}^{i}$. Using Lemma 5, it is sufficient to consider the case when all apprentices are trained in occupation $i$, that is, $n_{a}^{i}>0$ and $n_{a}^{j}=0 \forall j \neq i$.

## D.4. Dynamic Frictions

In this section, we consider a two-period version of the model including dynamic frictions. In period $t$, as in the baseline model, firms hire workers to produce and to train apprentices. In period $t+1$, firms can use both workers and previously trained apprentices in production.

With competitive labor markets, trained apprentices move freely after training, yielding the same results as in the baseline model. If the cost of training apprentices in period $t$ is larger than the cost of hiring workers $\left(w_{t}<\frac{w_{t} t_{a}+w_{a}}{\zeta_{a}}\right)$, firms choose not to train any apprentices; while in the reverse case $w_{t}>\frac{w_{t} t_{a}+w_{a}}{\zeta_{a}}$, they train as many apprentices as possible, $n_{a, t}^{*}=\left(\frac{p_{t} \gamma_{a}^{\gamma}}{w_{t} t_{a}+w_{a}}\right)^{1 /(1-\gamma)} z, n_{t}^{*}=t_{a} n_{a, t}^{*}$.

Now consider labor market frictions that prevent a fraction $\rho>0$ of apprentices from moving across firms after training. Firms can retain these apprentices at a discounted wage rate $w_{a}^{r}$ below their marginal product. In other words, there is wage compression. Firms choose $n_{t}$ workers and $n_{a, t}$ apprentices in period $t$, and $n_{a}^{r}$ apprentices to continue in period $t+1$ as well as $n_{t+1}$ workers to solve

$$
\begin{aligned}
& \max _{n_{t}, n_{a, t}, n_{a}^{r}, n_{t+1}} p_{t} f\left(n_{t}-t_{a} n_{a, t}, \zeta_{a} n_{a, t} ; z\right)+\beta p_{t+1} f\left(\zeta_{a}^{r} n_{a}^{r}+n_{t+1} ; z\right)-w_{t} n_{t}-w_{a} n_{a, t} \\
& \quad-\beta w_{t+1} n_{t+1}-\beta w_{a}^{r} n_{a}^{r} \\
& \text { s.t. } n_{t} \geq t_{a} n_{a, t}, n_{a, t} \geq 0, n_{a}^{r} \geq 0, n_{t+1} \geq 0, \rho n_{a, t} \geq n_{a}^{r}
\end{aligned}
$$

where $\zeta_{a}$ denotes the productivity of apprentices in period $t, \zeta_{a}^{r}$ the productivity of retained apprentices in $t+1$ and $\beta \in[0,1]$ the intertemporal discount rate.

We solve the model as in the baseline case, that is, $f\left(n_{t}-t_{a} n_{a t}, \zeta_{a} ; z\right)=z^{1-\gamma}\left(n_{t}-t_{a} n_{a t}+\right.$ $\left.\zeta_{a} n_{a t}\right)^{\gamma}$ and $f\left(\zeta_{a}^{r} n_{a}^{r}+n_{t+1} ; z\right)=\left(\zeta_{a}^{r} n_{a}^{r}+n_{t+1}\right)^{\gamma}$, but consider the additional restrictions. Again, under linear labor inputs, there are corner solutions. In this case, the FOCs of the
firm problem imply that with wage compression $w_{a}^{r}<\zeta_{a}^{r} w_{t+1}$, firms train apprentices if and only if

$$
\begin{equation*}
\underbrace{w_{t}}_{\text {Cost of hiring workers in } t} \geq \underbrace{\frac{w_{a}+t_{a} w_{t}}{\zeta_{a}}}_{\text {Cost of training apprentices in } t}-\underbrace{\frac{\rho \beta}{\zeta_{a}}\left(\zeta_{a}^{r} w_{t+1}-w_{a}^{r}\right)}_{t+1 \text { benefit }} \tag{8}
\end{equation*}
$$

Note that equation (8) generalizes the inequality from the frictionless case by including the additional benefits from training. On the other hand, if there is no wage compression such that $w_{a}^{r} \geq \zeta_{a}^{r} w_{t+1}$, the condition for firms to train apprentices is the same as in the baseline case. Moreover, if $t_{a} \geq \rho^{1-\gamma} \zeta_{a}^{\gamma}\left(\zeta_{a}^{r}\right)^{1-\gamma} \frac{p_{t}}{p_{t+1}} \frac{w_{a}^{r}}{w_{t}}-\frac{w_{a}}{w_{t}}$, firms also adjust the intensive margin, training more apprentices than in the frictionless case: $n_{a, t}^{f}=$ $\left(\frac{p_{t} \gamma s_{a}^{\gamma}}{w_{a}+t_{a} w_{t}-\beta \rho\left(\zeta_{a}^{w} w_{t+1}-w_{a}^{r}\right)}\right)^{\frac{1}{1-\gamma}} z>n_{a, t}^{*}$. Lemma 6 compiles these results, showing that frictions together with wage compression increase firms' willingness to train apprentices in the spirit of Acemoglu and Pischke $(1998,1999)$.

LEMMA 6: With frictions and wage compression, that is, $\rho>0$ and $\frac{w_{a}^{r}}{\zeta a}>w_{t+1}$, the number of trained apprentices increases relative to the baseline model.

For details of the calibration of this model extension, see Appendix E.8.

## D.5. Hiring Costs

In this section, we describe an extension of the model that allows for fixed and variable hiring costs $\kappa_{0}^{h}, \kappa_{1}^{h} \geq 0$. We suppose that firms incur these costs if they increase their size $n$ relative to their pre-regulation size $n^{*}$.

Formally, firm $\left(z, t_{a}\right)$ solves

$$
\begin{aligned}
& \max _{n, n_{a}, d_{f}} p f\left(n-t_{a} n_{a}, \zeta_{a} n_{a} ; z\right)-w n-w_{a} n_{a}-d_{f} \mathcal{F}_{a}\left(n, n_{a}\right) \\
& \quad-\underbrace{\left(\kappa_{0}^{h}+\kappa_{1}^{h} n\right) \mathbb{1}\left(n>n^{*}\right)}_{\text {Hiring Costs }} \quad \text { s.t. } t_{a} n_{a} \leq n, \\
& \left(n, n_{a}, d_{f}\right) \in \bigcup_{j}\left[N_{j-1}, N_{j}\right) \times\left[\underline{n}_{a}^{j}, \bar{n}_{a}^{j}\right] \times\{0\} \quad \text { or } \\
& \left(n, n_{a}, d_{f}\right) \in \bigcup_{j}\left[N_{j-1}, N_{j}\right) \times\left[0, \underline{n}_{a}^{j}\right] \times\{1\}, \quad \mathcal{F}_{a}\left(n, n_{a}\right)=\phi_{a}\left(\underline{n}_{a}^{j}-n_{a}\right)^{+},
\end{aligned}
$$

where $n^{*}\left(z, t_{a}\right)$ is the solution to the problem without regulation. We solve this model numerically, following an analogous procedure to in the baseline case. See Appendix E. 8 for details of the calibration.

## APPENDIX E: DETAILS of the Quantitative Exercises

## E.1. Parametric Fit of the Firm Size Distribution

Figure E1 shows the fit of various parametric distributions to prereform data. The Generalized Extreme Value distribution provides the best fit out of two- and three-parameter distributions commonly used to model productivity.


Figure E1.-Fitting the Prereform Firm Size Distribution. Notes: The figure shows the empirical firm size density pre-reform (1995 to 2002), as well as various parametric distributions fitted via maximum likelihood.

## E.2. Production Function Estimation

Table EI shows estimated labor shares $\gamma^{k}$ using six different methodologies: our baseline regression with time and firm fixed effects (FE), a simple OLS regression (OLS), as well as the methods proposed by Olley and Pakes (1996) (OP), Levinsohn and Petrin (2003) (LP), Wooldridge (2009) (W), and Ackerberg, Caves, and Frazer (2015) (LP-ACF). For the last four specifications, we suppose the production function depends on capital $K$, full-time labor $l$ and other intermediate inputs $m$, which we approximate empirically using energy, water and intermediate product expenditures.

## E.3. Moment Weights and Robustness

Table EII details the weights on each group of moments in the estimation. For the bunching and missing mass moments, we use weights corresponding to the observed prereform fraction of firms at each bunching or missing mass point. For instance, for the first bunching point among high-skill sector firms we weight the bunching mass at 14 by the fraction of high-skill sector firms of size 14 in the prereform data, $h_{14}^{s}$. Additionally, we divide the missing mass moments by five (the width of the potential missing mass window we consider), in order to make the total weight on missing mass moments comparable to the weight on bunching moments. We use the same procedure for weighting the average number of apprentices by firm size and the fraction of firms paying the fee by firm size, using as weights the prereform fraction of firms $h_{j}$. Finally, we equally weight those four

TABLE EI
ESTIMATED LABOR SHARES.

|  | (1) | $(2)$ | $(3)$ | $(4)$ | $(5)$ | (6) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | FE (Baseline) | OLS | OP | LP | W | LP-ACF |
| High-Skill | 0.61 | 0.30 | 0.57 | 0.34 | 0.31 | 0.42 |
| Low-Skill | 0.58 | 0.34 | 0.43 | 0.23 | 0.20 | 0.28 |

Note: The table shows labor shares obtained from different estimation methods. Columns (3) to (6) are estimated using the Stata program prodest by Mollisi and Rovigatti (2017).

TABLE EII
BASELINE MOMENT WEIGHTS.

| Weight | Moment Description |
| :--- | :--- |
| $\omega_{j}^{k}=\frac{1}{4} \frac{1}{2} h_{b(j)}$ | Bunching mass points |
| $\omega_{j}^{k}=\frac{1}{4} \frac{1}{2} \frac{1}{5} h_{m(j)}$ | Missing mass points |
| $\omega_{j}^{k}=\frac{1}{4} h_{j}$ | Average number of apprentices by firm size |
| $\omega_{j}^{k}=\frac{1}{4} h_{j}$ | Fraction of firms paying the fee by firm size |
| $\omega_{j}^{k}=\frac{1}{4}$ | Fraction of firms choosing maximum number of apprentices prereform |

Note: The table shows the weights on moments used in the SMM estimation.
groups of moments. Thus, the fraction of firms that choosing the maximum number of apprentices before the reform receive weight $\omega_{j}^{k}=\frac{1}{4}$.

As a robustness check, instead of equally weighting the four groups of moments, we weight them by the inverse of their variance obtained from 1000 bootstrap samples. Table EIII shows that this procedure results in somewhat extreme weights as the variance of some moments is small. For instance, virtually no high-skill sector firms choose the maximum number of apprentices before the reform, resulting in a small variance and a large weight on this moment. Similarly, very few low-skill sector firms pay the fee, attracting a large weight.

Figure E2 and Table EIV show that uniform weighting matches better the average number of apprentices (Fact 2) and the fraction paying the fee (Fact 3) for high-skill sectors. In particular, using inverse variance weights leads to somewhat overestimating the apprentice intake in high-skill sectors. Figure E3 shows that the inverse variance estimation results in a very similar estimated training cost distribution for low-skill sectors, but somewhat smaller training costs for high-skill sectors. Consistently, Table EV shows that quantitative results for low-skill sectors remain almost identical, but the inverse variance weighting over-estimates labor substitution in high-skill sectors.

TABLE EIII
INVERSE VARIANCE MOMENT WEIGHTS.

| Moment Group | (1) <br> High-skill | (2) <br> Low-Skill |
| :--- | :---: | :---: |
| Fraction choosing maximum apprentices pre-reform | 0.931 | 0.388 |
| Bunching and missing mass points | 0.065 | 0.100 |
| Average number of apprentices by firm size | 0.0001 | 0.0001 |
| Fraction paying the fee by firm size | 0.005 | 0.511 |

Note: The table shows weights on each group moments based on the inverse of their variance, obtained from 1000 bootstrap samples.


Figure E2.-Alternative Moment Weights and Model Fit. Notes: The figure depicts the model fit to targeted moments using the inverse variance weights from Table EIII. Panels (a) and (b) show the distribution of firm size (number of full-time workers) for prereform (1995-2002) and panels (c) and (d) show the firm size distribution post-reform (2003-2009). Panel (e) shows the number of apprentices by firm size, and panel (f) shows the fraction of firms paying the fee by firm size, both in the post-reform period.

TABLE EIV
COMPARING FIT ACROSS ESTIMATION PROCEDURES.

|  | Targeted Moments Error |  |  |  | (5) <br> Total Score |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Max. Apprentices Prereform | Bunching (Fact 1) | Apprentices (Fact 2) |  | $\sum \omega \frac{\mid \text { model-datal }}{0.5 \mid \text { model }\|+0.5\| \text { data } \mid}$ |
|  | A. Baseline |  |  |  |  |
| High-Skill Sectors | 0.000 | 0.137 | 0.105 | 0.060 | 0.302 |
| Low-Skill Sectors | 0.001 | 0.129 | 0.052 | 0.097 | 0.280 |
|  | B. Out-of-Sample Fit |  |  |  |  |
| High-Skill Sectors | 0.000 | 0.141 | 0.127 | 0.071 | 0.340 |
| Low-Skill Sectors | 0.001 | 0.132 | 0.051 | 0.057 | 0.240 |
|  | C. Inverse Variance Weights |  |  |  |  |
| High-Skill Sectors | 0.000 | 0.074 | 0.173 | 0.116 | 0.022 |
| Low-Skill Sectors | 0.001 | 0.129 | 0.052 | 0.097 | 0.037 |

Note: The table shows the model fit under the baseline estimation (panel A), the out-of-sample estimation (panel B), and the inverse variance weighting estimation (panel C). Columns (1) to (4) show the estimation error of the targeted moments $\frac{\mid \text { model-data } \mid}{0.5 \mid \text { model }|+0.5| \text { data }}$. Column (5) shows the total score function using uniform weights in panels A and B , and inverse variance weights from Table EIII in panel C.

## TABLE EV

Robustness of quantitative effects using alternative moment weights.

|  | (1) | (2) Aggregat | (3) Outcomes | (4) | Changes in Agents' Welfare $\Delta U_{j} / U^{*}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Workers | \% Workers | Apprentices | \% Output | Apprentices | Workers | Firms | Total |
|  | A. Uniform Weights |  |  |  |  |  |  |  |
| High-Skill Sectors | -1905 | -0.89 | 4937 | -0.34 | 0.27 | -0.54 | -0.46 | $-0.74$ |
| Low-Skill Sectors | -3519 | -1.67 | 17,866 | 0.30 | 1.13 | -0.97 | 0.13 | 0.30 |
| Total | -5423 | -1.28 | 22,803 | -0.06 | 0.65 | -0.73 | -0.20 | $-0.28$ |
|  |  |  | B. Inv | erse Variance | Weights |  |  |  |
| High-Skill Sectors | -2433 | -1.14 | 7492 | -0.37 | 0.41 | -0.70 | -0.35 | -0.64 |
| Low-Skill Sectors | -3493 | -1.66 | 17,869 | 0.31 | 1.13 | -0.96 | 0.13 | 0.31 |
| Total | -5926 | -1.40 | 25,361 | -0.07 | 0.73 | -0.81 | -0.14 | -0.22 |

[^3]

Figure E3.-Alternative Moment Weights and Training Costs. Notes: The figure shows estimated training cost distributions under the baseline weighting procedure (solid lines) and the alternative inverse variance weighting procedure (dashed lines). We normalize $\zeta_{a}=1$ as in main text Figure 7.

## E.4. Out-of-Sample Fit

In this section, we present results from an out-of-sample exercise. We randomly split the data set in half for all years and sectors, estimate the model using one-half and evaluate its fit using the other half. Table EIV shows the comparison of estimation errors of targeted moments as well as total score functions. Figure E4 depicts the out-of-sample model fit graphically.

## E.5. Truncated Normal and Uniform Training Cost Distributions

Figure E5 shows the fit of the estimated model assuming a truncated normal training cost distribution, and Figure E6 shows the fit assuming a uniform training cost distribution.

## E.6. Transferability of Training Skills

Figure E7 shows sensitivity analysis with respect to the skills transferability parameter $\chi_{a}^{k}$. In particular, we show how moving from $\chi_{a}=0$ to $\chi_{a}=1$ changes the effects on effective units of labor, output, profits, and total welfare.

## E.7. Effect of the Policy Decomposition on Aggregate Variables

Here, we show additional details of the policy decomposition described in Section 6.1 of the paper. Table EVI shows the effects of the different components of the regulation on aggregate outcomes. We consider four scenarios, (i) only apprentice quotas, (ii) only a reduction in apprentice minimum wages, (iii) combining quotas and the lower minimum wage, and (iv) the "full" regulation adding the possibility of paying the fee. We show that with only the minimum wage reduction, firms with low training costs substitute many of their workers for apprentices. Adding the quotas attenuate the displacement of workers by establishing a maximum on the number of apprentices. The minimum quota, on the other hand, mandates firms to train, which increases training in high-skill sectors. Finally, the possibility of paying the fee dampens the negative effects for those firms with very high training costs.


Figure E4.-Out-of-Sample Fit. Notes: The figure depicts the fit of the model estimated on half of the data to the other, untargeted half of the data. Panels (a) and (b) show the distribution of firm size (number of ful-l-time workers) for pre-reform (1995-2002) and panels (c) and (d) show the firm size distribution post-reform (2003-2009). Panel (e) shows the number of apprentices by firm size, and panel (f) shows the fraction of firms paying the fee by firm size, both in the post-reform period.


Figure E5.-Model Fit Under Truncated Normal Training Costs. Notes: The figure depicts the fit of the model under a truncated normal training cost distribution. Panels (a) and (b) show the estimated distributions of net training costs and apprentices' marginal productivity, respectively. Panels (c) and (d) show the firm size distribution post-reform (2003-2009). Panel (e) shows the number of apprentices by firm size, and panel (f) shows the fraction of firms paying the fee by firm size, both in the post-reform period.


Figure E6.-Model Fit Under Uniform Training Cost Distribution. Notes: The figure depicts the fit of the model under a uniform training cost distribution. Panels (a) and (b) show the estimated distributions of net training costs and apprentices' marginal productivity, respectively. Panels (c) and (d) show the firm size distribution post-reform (2003-2009). Panel (e) shows the number of apprentices by firm size, and panel (f) shows the fraction of firms paying the fee by firm size, both in the post-reform period.


Figure E7.-Sensitivity Analysis: Skills Transferability. Notes: The figure shows dynamic effects of the regulation under different values of the skill transferability parameter $\chi_{a}$. Effects are shown on effective units of labor (panel a), output (panel b), profits (panel c), and total welfare (panel d).

## E.8. Calibration of Extensions and Comparative Statics

Dynamic Frictions. We use PILA data to calibrate variables related to the dynamic frictions model from Section D.4. Using the information from Table BIV, we calibrate the probability of apprentices staying in the same firm in the month after graduation (conditional on staying in the sample) to $\rho^{s}=0.74$ and $\rho^{u}=0.75$, respectively, in high-skill and low-skill sectors. These apprentices earn average wages of $w_{a}^{r s}=1.47$ and $w_{a}^{r, u}=1.35$. Finally, we use the average observed wages of all other workers by sector to approximate $w_{t}^{s}=w_{t+1}^{s}=3.15$ and $w_{t}^{u}=w_{t+1}^{u}=2.42$. To obtain an upper bound of additional benefits from training, we set $\beta=1$ and $\zeta_{a}^{r, k}=1, \forall k$. Using these values, we compute the additional benefits in each sector and adjust the estimated training cost distribution (see Figure E8). Table EVII shows quantitative effects of the regulation with and without dynamic frictions. Figures E9 and E10 show comparative statics with respect to $\rho$ and $w_{a}^{r}$, respectively.

TABLE EVI
Policy decomposition: Aggregate variables.

|  | (1) <br> Workers | (2) \% Workers | (3) <br> Apprentices | (4) <br> \% Output |
| :---: | :---: | :---: | :---: | :---: |
| A. Only Quotas |  |  |  |  |
| High-Skill | -9258 | -4.34 | 8891 | -3.17 |
| Low-Skill | -3859 | -1.83 | 9514 | -0.42 |
| Total | -13,118 | -3.09 | 18,405 | -1.95 |
| B. Only $\downarrow w_{a}$ |  |  |  |  |
| High-Skill | -2345 | -1.10 | 14,205 | 0.16 |
| Low-Skill | -28,053 | -13.30 | 156,683 | 3.86 |
| Total | -30,398 | -7.16 | 170,887 | 1.80 |
| C. Quotas $+\downarrow w_{a}$ |  |  |  |  |
| High-Skill | -8024 | -3.76 | 9786 | -2.76 |
| Low-Skill | -3519 | -1.67 | 17,866 | 0.30 |
| Total | -11,543 | -2.72 | 27,653 | -1.41 |
| D. Full Regulation |  |  |  |  |
| High-Skill | -1905 | -0.89 | 4937 | -0.34 |
| Low-Skill | -3519 | -1.67 | 17,866 | 0.30 |
| Total | -5423 | -1.28 | 22,803 | -0.06 |

Note: The table shows the effects of apprenticeship regulation on aggregate outcomes, namely on the number of workers, the number of apprentices, and output. Effects are shown under four scenarios, namely only apprentice quotas (panel A), only a lower apprentice minimum wage (panel B), quotas plus the lower minimum wage (panel C), and the full regulation featuring quotas, the lower minimum wage, and the possibility of paying the fee (panel D)


Figure E8.-Dynamic Frictions and Estimated Training Costs. Notes: The figure shows estimated training cost distributions from the baseline model (solid lines) and under dynamic frictions (dashed lines). We normalize $\zeta_{a}=1$ as in main text Figure 7.

TABLE EVII
QUANTITATIVE EFFECTS WITH AND WITHOUT DYNAMIC FRICTIONS.


Note: Columns (1) to (4) of the table show the effects of the apprenticeship regulation on aggregate outcomes, namely on the number of workers, the number of apprentices, and output. Columns (5) to (7) show the effect on the welfare of apprentices, workers, and firm owners. Column (8) shows the sum of welfare effects across the groups of agents from columns (5) to (7). Panel A shows effects under the baseline model, and panel B shows effects with dynamic frictions.

Hiring Costs. We reestimate the training cost distribution under the model with hiring costs from Section D.5, setting the cost parameters to: fixed hiring costs $\kappa_{0}^{h}=0.1$ and variable hiring costs $\kappa_{1}^{h}=0.05$. These parameters are chosen to be large enough to have some effects on the estimated training cost distribution (see Figure E11), without radically changing the model fit (see Table EVIII). Choosing larger values of the hiring cost parameters would increasingly worsen the fit of the estimated model.

## E.9. Additional Results on Counterfactual Exercises

## Subsidizing Apprenticeship Training

The linearity of the model implies corner solutions. Lemma 7 characterizes this solution.

Lemma 7: Suppose a firm $\left(t_{a}, z\right)$ solves (5):
i. If $(1-\varsigma) w_{a}+w(1+\tau) t_{a}>w(1+\tau)$, then the firm avoids apprentices and chooses

$$
n^{*}=\left(\frac{p \gamma}{w(1+\tau)}\right)^{1 /(1-\gamma)} z, \quad n_{a}^{*}=0
$$

TABLE EVIII
Hiring costs, training costs, and model fit.

|  | $(1)$ <br> Baseline | (2) <br> Fixed Hiring Cost $\kappa_{0}^{h}=0.1$ | (3) <br> Variable Hiring Cost $\kappa_{1}^{h}=0.05$ <br> Average $t_{a}$ in High-Skill <br> 1.188 <br> 1.210 <br> Average $t_{a}$ i Low-Skill <br> Total Score SMM 0.753 |
| :--- | :---: | :---: | :---: |

[^4]

Figure E9.-Comparative Statics: Frictions. Notes: The figure shows comparative statics with respect to the probability of staying in the firm ( $\rho$ ), based on estimating the dynamic model described in Section D.4.
ii. If $(1-\varsigma) w_{a}+w(1+\tau) t_{a}<w(1+\tau)$, then the firm trains apprentices and chooses

$$
n^{*}=t_{a} n_{a}^{*}, \quad n_{a}^{*}=\left(\frac{p \gamma \zeta_{a}^{\gamma}}{\tilde{w}}\right)^{1 /(1-\gamma)} z, \quad \text { where } \tilde{w}:=w_{a}(1-\varsigma)+w(1+\tau) t_{a}
$$

Case (i) describes the situation where the cost of training apprentices is higher than the total cost of hiring workers, once the tax and the subsidy are taken into account. The tax scheme harms firms with high $t_{a}$, since these training costs are not directly covered by the subsidy. On the other hand, in Case (ii), when the tax and the subsidy are high enough, firms are incentivized to train. Thus, the demand for apprentices depends on whether the subsidy covers all monetary training expenses.


Figure E10.-Comparative Statics: Wage Compression. Note: The figure shows the comparative statics with respect to the wages of retained apprentices $\left(w_{a}^{r}\right)$, based on estimating the dynamic model described in Section D.4.

Additionally, we ensure the policy is budget-balanced. Total revenue from payroll taxes is $\operatorname{Rev}(\tau, \varsigma):=\sum_{k} \tau w^{k} N^{k}(\tau, \varsigma)$. Total subsidy payments to firms equal

$$
\operatorname{Sub}(\tau, \varsigma):=\sum_{k} F^{k} \iint \mathcal{S}^{*}\left(t_{a}, z ; \tau, \sigma\right) d \mathcal{Z}^{k}(z) d \mathcal{T}^{k}\left(t_{a}\right)
$$

where $F^{k}$ is the number of firms in sector $k$ and $\mathcal{S}^{*}\left(t_{a}, z ; \tau, \sigma\right)$ denotes the amount of subsidy optimally taken by firm $\left(t_{a}, z\right)$ when facing policy $(\tau, \sigma)$. In Section 6 of the paper, we consider the set of budget-balanced policies $(\tau, \varsigma)$ such that $\operatorname{Rev}(\tau, \varsigma)-\operatorname{Sub}(\tau, \varsigma)=0$.


Figure E11.-Hiring Costs and Estimated Training Costs. Notes: The figure shows estimated training cost distributions from the baseline model (solid lines in both panels), from the model with with fixed hiring costs (dashed lines in panel a), and from the model with variable hiring costs (dashed lines in panel b). We normalize $\zeta_{a}=1$ as in main text Figure 7.

## Sector-Specific Apprentice Minimum Wage

In the following, we describe the details of computing the sector-specific minimum wage counterfactual policy. Let $n_{a}^{k, \min }\left(z, t_{a} ; w_{a}^{k}, w^{k}\right)$ denote the solution to the firm maximization problem when the wage of workers is $w^{k}$ and the apprentice wage is $w_{a}^{k}$. For each sector $k$, we compute the minimum wage for apprentices $w_{a}^{* k}$ that solves

$$
\begin{equation*}
N_{a}^{* k}=F^{k} \iint n_{a}^{k, \min }\left(z, t_{a} ; w_{a}^{* k}, w^{k}\right) d \mathcal{Z}^{k}(z) d \mathcal{T}^{k}\left(t_{a}\right) \tag{9}
\end{equation*}
$$

where $N_{a}^{* k}$ denotes the total number of apprentices trained in sector $k$ under the original regulation. We take worker wages and structural parameter estimates from the baseline model.

Under linear labor inputs, a firm $\left(t_{a}, z\right)$ trains apprentices if the total cost of training is smaller than that of hiring workers, and chooses $n_{a}^{k, \text { min }}=\left(\frac{p^{k} \gamma^{k}\left(\xi_{a}^{k} \gamma^{k}\right.}{w_{a}^{k}+w^{k} t_{a}}\right)^{1 /\left(1-\gamma^{k}\right)} z$ and $n^{k, \text { min }}=t_{a} n_{a}^{k, \text { min }}$. Conversely, the firm does not train apprentices if $\frac{w_{a}^{k}+w^{k} t_{a}}{\zeta_{a}^{k}}>w^{k}$, such that $n_{a}^{k, \text { min }}=0$ and $n^{k, \text { min }}=\left(\frac{p^{k} \gamma^{k}}{w^{k}}\right)^{1 /\left(1-\gamma^{k}\right)} z$. We solve equation (9) numerically and obtain apprentice minimum wages of $w_{a}^{* s}=0.77$ in high-skill sectors and of $w_{a}^{* u}=0.98$ in low-skill sectors.

In Figure E12, we target the same total number of apprentices as under the baseline regulation, but change the allocation of apprentices across sectors. We show that the effect of changing the fraction of apprentices trained in high-skill sectors on the number of apprentices per worker, number of workers, total output, and total welfare.


Figure E12.-Varying the Share of Training in High-Skill Sectors. Notes: The figure shows the effects of varying the share of training in high-skill sectors by varying sector-specific minimum wages, holding the total number of apprentices fixed. Panel (a) shows the number of apprentices per worker, panel (b) shows the percentage change in the number of workers, panel (c) shows the percentage change in aggregate output, and panel (d) the change in total welfare measured as the change in total aggregate utility relative to prereform utility, $\Delta \mathcal{U} / \mathcal{U}^{*}$.

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[^1]:    Note: The table shows characteristics and indicators for the type of response to the apprenticeship regulation for three groups of firms: firms that bunch at (above) regulation thresholds (column 1), firms that bunch below regulation thresholds (column 2), and all firms (column 3). Only observations in the post-reform years 2003 to 2009 are included.

[^2]:    Note: The table shows information on apprentices' firm and sector transitions based on the PILA data. Panel A shows the probability of an apprentice remaining observed in the data. Panel B reports the transition probabilities and wages of apprentices moving across firms and industries during the sample period. Wages are shown in units relative to the minimum wage with standard deviations in parentheses.

[^3]:    Note: Columns (1) to (4) of the table show the effects of the apprenticeship regulation on aggregate outcomes, namely on the number of workers, the number of apprentices, and output. Columns (5) to (7) show the effect on the welfare of apprentices, workers, and firm owners. Column (8) shows the sum of welfare effects across the groups of agents from columns (5) to (7). Panel A shows results under the baseline uniform moment weights, and panel B shows results under the alternative inverse variance weights.

[^4]:    Note: The table shows average training costs by sector and the model fit (measured by the score function of the SMM estimation) under the baseline model (column 1), the model with fixed hiring costs (column 2) and the model with variable hiring costs (column 3).

